

### 17.1.1 Two Restricted Likelihoods That Can't Be Re-expressed

The first example is an additive model with two predictors,  $x_{1i}$  and  $x_{2i}$ , for the  $i^{\text{th}}$  observation. Given a spline basis for each predictor, Chapter 4 showed how this model can be written in mixed-linear-model form as

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}_1\mathbf{u}_1 + \mathbf{Z}_2\mathbf{u}_2 + \boldsymbol{\varepsilon}, \quad (17.1)$$

where  $\mathbf{X}\boldsymbol{\beta}$  includes the unshrunk parts of the splines for both predictors and  $\mathbf{Z}_j\mathbf{u}_j$  contains the shrunk part of the spline for predictor  $j$ , so the random effects design matrix in standard form is  $\mathbf{Z} = (\mathbf{Z}_1|\mathbf{Z}_2)$ . If the spline for predictor  $j$  has  $K_j$  knots, then the standard-form random effect  $\mathbf{u} = (\mathbf{u}'_1, \mathbf{u}'_2)'$  has covariance matrix

$$\mathbf{G} = \begin{bmatrix} \sigma_{s_1}^2 \mathbf{I}_{K_1} & \mathbf{0} \\ \mathbf{0} & \sigma_{s_2}^2 \mathbf{I}_{K_2} \end{bmatrix}. \quad (17.2)$$

To derive the restricted likelihood, let  $\mathbf{P}_X^c = \mathbf{I}_n - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}$  be the projection onto the orthogonal complement of the column space of the fixed effect design matrix  $\mathbf{X}$ . If  $p$  is the rank of  $\mathbf{X}$ , let  $\mathbf{H}$  be an  $n \times (n-p)$  matrix with orthonormal columns such that  $\mathbf{H}\mathbf{H}' = \mathbf{P}_X^c$ , e.g.,  $\mathbf{H}$  has columns equal to the eigenvectors of  $\mathbf{P}_X^c$  with eigenvalue 1. Pre-multiply equation (17.1) by  $\mathbf{H}'$ , so that

$$\mathbf{w} = \mathbf{H}'\mathbf{y} = \mathbf{H}'\mathbf{Z}_1\mathbf{u}_1 + \mathbf{H}'\mathbf{Z}_2\mathbf{u}_2 + \boldsymbol{\phi}. \quad (17.3)$$

Under the usual assumption that  $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma_e^2 \mathbf{I}_n)$ ,  $\boldsymbol{\phi} = \mathbf{H}'\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma_e^2 \mathbf{I}_{n-p})$ . The restricted likelihood is the likelihood for  $(\sigma_{s_1}^2, \sigma_{s_2}^2, \sigma_e^2)$  arising from equation (17.3), so the log restricted likelihood is

$$\begin{aligned} & -0.5 \log |\sigma_{s_1}^2 \mathbf{H}'\mathbf{Z}_1\mathbf{Z}'_1\mathbf{H} + \sigma_{s_2}^2 \mathbf{H}'\mathbf{Z}_2\mathbf{Z}'_2\mathbf{H} + \sigma_e^2 \mathbf{I}_{n-p}| \\ & -0.5 \mathbf{w}' (\sigma_{s_1}^2 \mathbf{H}'\mathbf{Z}_1\mathbf{Z}'_1\mathbf{H} + \sigma_{s_2}^2 \mathbf{H}'\mathbf{Z}_2\mathbf{Z}'_2\mathbf{H} + \sigma_e^2 \mathbf{I}_{n-p})^{-1} \mathbf{w}. \end{aligned} \quad (17.4)$$

To express this restricted likelihood in the desired scalar form, we need to diagonalize three matrices simultaneously:  $\mathbf{H}'\mathbf{Z}_1\mathbf{Z}'_1\mathbf{H}$ ,  $\mathbf{H}'\mathbf{Z}_2\mathbf{Z}'_2\mathbf{H}$ , and  $\mathbf{I}_{n-p}$ . To do that, we need to find an orthogonal matrix  $\Gamma$  such that  $\Gamma'\mathbf{H}'\mathbf{Z}_1\mathbf{Z}'_1\mathbf{H}\Gamma$  and  $\Gamma'\mathbf{H}'\mathbf{Z}_2\mathbf{Z}'_2\mathbf{H}\Gamma$  are both diagonal. Such a  $\Gamma$  exists if and only if  $\mathbf{H}'\mathbf{Z}_1\mathbf{Z}'_1\mathbf{H}\mathbf{H}'\mathbf{Z}_2\mathbf{Z}'_2\mathbf{H}$  is symmetric (Graybill 1983, Theorem 12.2.12), which is not true in general. An exercise has you check this.

To ensure that this restricted likelihood has the desired scalar form, a sufficient condition is a form of balance in which the two predictors  $x_{1i}$  and  $x_{2i}$  are observed on a rectangular grid with the same number of replicates at each grid point. I suspect this is also a necessary condition but I have not proved it. An exercise examines this question.

The second example of a restricted likelihood that can't be re-expressed is the ICAR model with two classes of neighbor relations (2NRCAR) introduced in Section 14.2. (This proof is taken from Reich et al. 2007.) The  $n$ -vector  $\mathbf{y}$  is modeled as  $\mathbf{y} = \boldsymbol{\delta} + \boldsymbol{\varepsilon}$ , where  $\boldsymbol{\varepsilon} \sim N_n(\mathbf{0}, \sigma_e^2 \mathbf{I}_n)$  and  $\boldsymbol{\delta}$  has an improper normal density specified in terms of its precision matrix:

$$f(\boldsymbol{\delta} | \sigma_{s_1}^2, \sigma_{s_2}^2) \propto \exp \left[ -\frac{1}{2} \boldsymbol{\delta}' (\mathbf{Q}_1 / \sigma_{s_1}^2 + \mathbf{Q}_2 / \sigma_{s_2}^2) \boldsymbol{\delta} \right]; \quad (17.5)$$