

Linear Factor Analysis - The Model

$$\mathbf{x} = \boldsymbol{\mu} + \boldsymbol{\Lambda}\mathbf{f} + \boldsymbol{\epsilon} \quad (1)$$

\mathbf{x} : p -dimensional vector of continuous observed variables

\mathbf{f} : q -dimensional vector of underlying latent factors. Often called “common factors”. Assume \mathbf{f} is random such that $E(\mathbf{f}) = \mathbf{0}$ and $Var(\mathbf{f}) = \boldsymbol{\Phi}$

$\boldsymbol{\epsilon}$: p -dimensional vector of random error. Often called “unique factors” or “specific factors”. Assume $E(\boldsymbol{\epsilon}) = \mathbf{0}$ and $Var\boldsymbol{\epsilon} = \boldsymbol{\Psi}$. Element along the diagonal of $\boldsymbol{\Psi}$ often called “uniquenesses” or “specific variances”

$\boldsymbol{\Lambda}$: $p \times q$ matrix of scalars called “factor loadings”. This matrix describes how the observed variables \mathbf{x} are related to the latent factors \mathbf{f} .

$\boldsymbol{\mu}$: $p \times 1$ vector of scalars. Often ignored, most software assumes by default that $\boldsymbol{\mu} = \mathbf{0}$ and analyze centered \mathbf{x} variables, i.e. analyze $\mathbf{x} - \bar{\mathbf{x}}$

Slide 1

Linear Factor Analysis - The Model

$$\mathbf{x} = \boldsymbol{\mu} + \boldsymbol{\Lambda}\mathbf{f} + \boldsymbol{\epsilon}$$

An assumption is made for model (1) that $Cov(\mathbf{f}, \boldsymbol{\epsilon}) = \mathbf{0}$. This is a very critical assumption because what it means is that the variability in \mathbf{x} can be separated into two additive parts, one coming from the **common factors** and one coming from **specific factors** or the errors. That is there is no covariance between the two. Specifically,

$$\begin{aligned} Var(\mathbf{x}) &= \boldsymbol{\Lambda}Var(\mathbf{f})\boldsymbol{\Lambda}' + Var(\boldsymbol{\epsilon}) + 2Cov(\mathbf{f}, \boldsymbol{\epsilon}) \\ Var(\mathbf{x}) &= \boldsymbol{\Lambda}\boldsymbol{\Phi}\boldsymbol{\Lambda}' + \boldsymbol{\Psi}. \end{aligned}$$

It is assumed that equation (1) holds for each individual in the population and thus for $i = 1 \dots n$ independently sampled individuals we have

$$\mathbf{x}_i = \boldsymbol{\mu} + \boldsymbol{\Lambda}\mathbf{f}_i + \boldsymbol{\epsilon}_i.$$

Slide 2

Slide 3

As an example of how to write model (1), here are the equations when $p = 5$ and $q = 2$ and it is assumed that Ψ is a diagonal matrix, for $i = 1 \dots n$ we have:

$$\begin{aligned} x_{1i} &= \mu_1 + \lambda_{11}f_{1i} + \lambda_{12}f_{2i} + \epsilon_{1i} \\ x_{2i} &= \mu_2 + \lambda_{21}f_{1i} + \lambda_{22}f_{2i} + \epsilon_{2i} \\ x_{3i} &= \mu_3 + \lambda_{31}f_{1i} + \lambda_{32}f_{2i} + \epsilon_{3i} \\ x_{4i} &= \mu_4 + \lambda_{41}f_{1i} + \lambda_{42}f_{2i} + \epsilon_{4i} \\ x_{5i} &= \mu_5 + \lambda_{51}f_{1i} + \lambda_{52}f_{2i} + \epsilon_{5i} \end{aligned}$$

or in matrix/vector notation

$$\begin{pmatrix} x_{1i} \\ x_{2i} \\ x_{3i} \\ x_{4i} \\ x_{5i} \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \mu_4 \\ \mu_5 \end{pmatrix} + \begin{pmatrix} \lambda_{11} & \lambda_{12} \\ \lambda_{21} & \lambda_{22} \\ \lambda_{31} & \lambda_{32} \\ \lambda_{41} & \lambda_{42} \\ \lambda_{51} & \lambda_{52} \end{pmatrix} \begin{pmatrix} f_{1i} \\ f_{2i} \end{pmatrix} + \begin{pmatrix} \epsilon_{1i} \\ \epsilon_{2i} \\ \epsilon_{3i} \\ \epsilon_{4i} \\ \epsilon_{5i} \end{pmatrix}$$

and we also must specify the variance structure of \mathbf{f} and ϵ

$$Var(\mathbf{f}_i) = Var \begin{pmatrix} f_{1i} \\ f_{2i} \end{pmatrix} = \Phi = \begin{pmatrix} \phi_{11} & \phi_{12} \\ \phi_{12} & \phi_{22} \end{pmatrix}$$

$$Var(\epsilon_i) = Var \begin{pmatrix} \epsilon_{1i} \\ \epsilon_{2i} \\ \epsilon_{3i} \\ \epsilon_{4i} \\ \epsilon_{5i} \end{pmatrix} = \Psi = \begin{pmatrix} \psi_1 & 0 & 0 & 0 & 0 \\ 0 & \psi_2 & 0 & 0 & 0 \\ 0 & 0 & \psi_3 & 0 & 0 \\ 0 & 0 & 0 & \psi_4 & 0 \\ 0 & 0 & 0 & 0 & \psi_5 \end{pmatrix}$$

Communalities

- Communalities are the diagonal elements of $\Lambda\Lambda'$ when the factors are assumed to be uncorrelated
- Communalities are the part of the variance of each observed variable which is due to the q underlying factors

$$\begin{aligned} x_1 &= \mu_1 + \lambda_{11}f_1 + \lambda_{12}f_2 + \epsilon_1 \\ x_2 &= \mu_2 + \lambda_{21}f_1 + \lambda_{22}f_2 + \epsilon_2 \\ x_3 &= \mu_3 + \lambda_{31}f_1 + \lambda_{32}f_2 + \epsilon_3 \\ x_4 &= \mu_4 + \lambda_{41}f_1 + \lambda_{42}f_2 + \epsilon_4 \\ x_5 &= \mu_5 + \lambda_{51}f_1 + \lambda_{52}f_2 + \epsilon_5 \end{aligned}$$

$$\begin{aligned} Var(x_1) &= \lambda_{11}^2 + \lambda_{12}^2 + \psi_1 \\ Var(x_2) &= \lambda_{21}^2 + \lambda_{22}^2 + \psi_2 \\ Var(x_3) &= \lambda_{31}^2 + \lambda_{32}^2 + \psi_3 \\ Var(x_4) &= \lambda_{41}^2 + \lambda_{42}^2 + \psi_4 \\ Var(x_5) &= \lambda_{51}^2 + \lambda_{52}^2 + \psi_5 \end{aligned}$$

The communality for x_2 is $\lambda_{21}^2 + \lambda_{22}^2$. On the other hand, ψ_2 is the part of the variability in x_2 that is unique to x_2 , i.e. is not shared with other observed variables, "is not common".

Slide 4

Slide 5

Exploratory factor analysis vs. Confirmatory factor analysis (EFA vs. CFA)

EFA general purposes:

- To determine how many underlying factors are necessary to explain most of the correlations and variance in the data.
- To determine the relationship via **rotation** between each of these underlying factors with each of the observed variables in a meaningful way so that the factors can be interpreted and named.
- To weed out observed variables that do not tend to measure well the underlying factors shared by the other variables.
- To propose blocks of variables that may be subsequently be used to create a simple sum scale.
- To propose a CFA model

In EFA every element in Λ is estimated and it is assumed that Ψ is diagonal. Also, it is common to assume that $Var(\mathbf{f}) = \Phi = \mathbf{I}$, i.e. the factors are uncorrelated with variance 1 (but this is not a necessary assumption, it is dropped when examining oblique rotations).

Slide 6

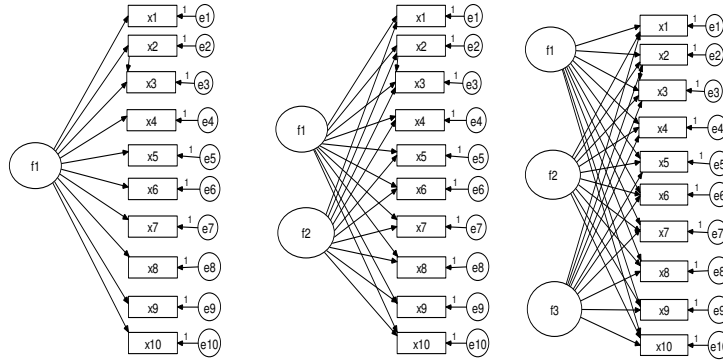
Exploratory factor analysis vs. Confirmatory factor analysis (EFA vs. CFA)

CFA general purposes:

- To define a measurement model for the relationship between multivariate observations and underlying factors
- To test the statistical significance of factor loadings and correlations. Note this testing cannot currently be done in the EFA model. Thus one may be interested in testing whether rotated factor loadings from an EFA that look "close to zero" are, in fact, significantly different from zero or not.
- To test whether the measurement model for one group is the same as the measurement model for some other group
- As a precursor to a Structural equation model

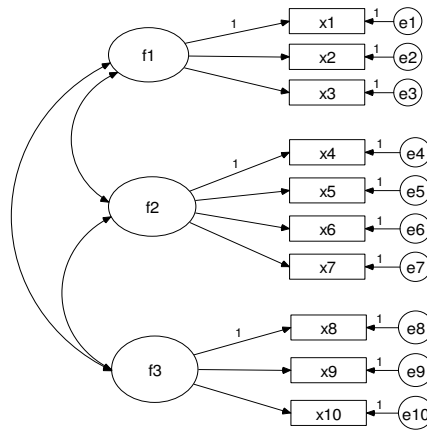
In CFA usually several elements in Λ are fixed to zero and it is possible to consider correlated ϵ which means that Ψ is not necessarily diagonal. Furthermore, it is usually assumed that the factors are correlated so that no restriction is placed on Φ .

Slide 7



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Slide 8



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Exploratory factor analysis vs. Confirmatory factor analysis (EFA vs. CFA)

What are the differences in the two models?

In the EFA model, there are arrows going from every factor to every observed variable. Also there are no double headed curved arrows going from factor to factor.

In the CFA model, there are clearly a lot less arrows going from factors to observed variable and in particular each observed variable has exactly 1 arrow hitting it - this is called simple structure. Also there are curved double headed arrows going from each factor to each other factor.

Look at papers:

- Hurley, A., Scandura, T, Sriesheim, C. Brannick, M., Seers, A, Vandenberg, R., and Williams, L. (1997) "Exploratory and confirmatory factor analysis: guidelines, issues, and alternatives" *Journal of organizational behavior*, 18, 667-683.
- Gerbing DW and Hamilton JG "Viability of Exploratory factor analysis as a precursor to confirmatory factor analysis", *Structural equation modeling*, 3(1), 62-72.

Slide 9

Outline of Exploratory Factor Analysis

- Factor extraction method
 - principal component analysis
Starts with correlation matrix and extracts factors (by calculating eigenvectors and eigenvalues) from it directly.
 - common factor analysis (Maximum Likelihood, Least squares, or Principal factor method)
Starts with correlation matrix but then makes a "reduced correlation matrix" meant to represent the correlation of the "common part", then extracts factors (by calculating eigenvectors and eigenvalues or using a discrepancy function) from the "reduced correlation matrix".
- Choose the number of factors
 - examine eigenvalues: % variance explained, eigenvalue > 1, scree plot
 - Model fit statistics: Chi-square value, RMSR root mean square residuals for correlation matrix, RMSEA (Cudeck) function of chi-square taking into account sample size, < .05 good fit, < .08 ok fit

Slide 10

Outline of Exploratory Factor Analysis - Continued

- Interpreting the factors - Rotation
 - rotate to simple structure - orthogonal (varimax) or oblique (promax) correlated factors
 - examining size of loadings (stat sig?), cross loadings
- Determining quality of factors and items
 - subject matter interpretation
 - number of items loading on the factor
 - variance explained by the factor
 - eliminate poor items and factors and repeat EFA
- Possible next steps
 - Verify structure using confirmatory factor analysis
 - Calculate factor score estimates $E(f|X)$, Mplus produces a “factor determinacy” value which represents correlation between factor and the estimated factor score
 - Calculate simple sum scores of items associated with each f with simple structure, use Cronbach alpha as estimate of reliability

Slide 11

Maximum number of factors possible to extract

- When the PCA method is used, it is possible to extract as many factors as there are variables.
- For the common factor methods which assume that the factors explain only the systematic part of the data (not the specific errors), the number of factors is constrained. The following must be true $\frac{p(p+1)}{2} - pq - p + \frac{q(q-1)}{2} \geq 0$. So for example when $p = 3$ or $p = 4$, q must be no greater than 1. When $p = 20$, q can be up to 14.

This restriction is related to the degrees of freedom associated with the common factor model. The number of parameters being estimated must be less than or equal to the amount of unconstrained data. For the EFA model the d.f. are $\frac{p(p+1)}{2} - pq - p + \frac{q(q-1)}{2}$. The first term is the number of unique elements in the sample covariance matrix \mathbf{S} , then we are going to estimate pq factor loadings in $\mathbf{\Lambda}$ and p variances in $\mathbf{\Psi}$ and since we have to put restrictions in order to fix rotation we get those $\frac{q(q-1)}{2}$ d.f. back. These $\frac{q(q-1)}{2}$ d.f. correspond to fixing the $\mathbf{\Lambda}'\mathbf{\Psi}^{-1}\mathbf{\Lambda}$ off diagonal elements being fixed to zero.

Slide 12

Rotation in EFA - this is an issue related to model identification

No matter what estimation procedure, for exploratory factor analysis we get estimates that look like:

$$\hat{\Sigma} = \hat{\Lambda}\hat{\Lambda}' + \hat{\Psi}$$

We can get the exact same $\hat{\Sigma}$ by taking

$$\hat{\Sigma} = \hat{\Lambda}TT'\hat{\Lambda}' + \hat{\Psi}$$

where T is a $q \times q$ matrix such that $TT' = \text{Identity matrix}$.

Thus

$$\hat{\Sigma} = \hat{\Lambda}^*\hat{\Lambda}^{*'} + \hat{\Psi}$$

where $\hat{\Lambda}^*$ is the **orthogonally rotated** factor loading matrix, i.e. $\hat{\Lambda}^* = \hat{\Lambda}T$. BUT, There are an infinite number of matrices T that satisfy $TT' = I$.

Slide 13

Rotation

- Why do we bother to rotate? – We want to find out which variables “stick together”
- How to choose a T ? – Find one that gives close to “simple” structure to the factor loadings
- “Simple structure” means that each observed variable only loads on one factor. In an AMOS notation, each observed variable only has one arrow going to it from a factor. Simple structure allows you to say, for example, “observed variable 1 is a direct measure of only factor 1”, instead of “variable 1 is direct measure of factor 1, 2, and 3”
- What about if simple structure is not found by orthogonal rotation?
- NOTE: implicitly orthogonal rotation keeps the assumption that $Var(\mathbf{f}) = I$, i.e. the factors are uncorrelated.
- Maybe the assumption that $Var(\mathbf{f}) = I$ should be dropped. This leads to **oblique rotations** of the factor loadings.
- We can search for

$$\hat{\Sigma} = \hat{\Lambda}^{**}\Phi\hat{\Lambda}^{**'} + \hat{\Psi}$$

so that $\hat{\Lambda}^{**}$ has “simple structure” and $\Phi \neq I$

Slide 14

Matching items to factors - Examine the rotated loadings

Throughout it is necessary to keep in mind whether the items loading on a given factor share some conceptual meaning.

- What is a BIG loading?
 - Common rule of thumb, if the absolute value of the standardized loading is $> .3$, the variable is relevant for the particular factor.
 - In 1994 Cudeck, R. and O'Dell, L. "Applications of standard error estimates in unrestricted factor analysis: Significance Tests for factor loadings and correlations" *Psychological Bulletin* argue that the $> .3$ cutoff is too simplistic and encourage sig testing.
 - Now possible to get standard errors from SAS Proc factor when using method = ml and the se option.
 - Additional work on calculated standard errors is Ogasawara (1998) "Standard errors for rotation matrices with an application to the promax solution". And there is a specialized program available from Michael Browne's website (at Ohio State University) called CEFA that can calculate these standard errors. All of these results are based on asymptotic results, thus it is still not clear how useful they are for problems with small sample sizes.

Slide 15

Matching items to factors - Continued

- Examine the communality. Rule of thumb, if the communality is less than $.10$ then the observed variable may be deleted. A communality $< .10$ means that less than 10% of the variability in the observed variable is explained by all of the common factors.
- Items with strong loadings on multiple factors (4 options)
 - allow the item to load on more than one factor
 - choose the item to load on the factor it has the largest loading with
 - choose the item to load on the factor it conceptually goes best with
 - drop the item
- Cross Validation. If the sample size is large enough, an intuitive way to gain an idea about the stability of the factor structure is to split the sample randomly into two equal parts and then fit the model to each part. If the factor structure is similar, this tends to increase our confidence in the genuineness of the factors.

Slide 16

Variance explained by factors

In EFA the sum of the total variance explained by k factors is same for any orthogonal rotation of the k factors. The difference is how the variance is spread out across the factors.

On the other hand, when an oblique rotation is used, summarizing the variance explained by the factors is tricky because we need to talk about the variance explained after adjusting for the other factors (since they are correlated).

Slide 17

School Subjects example from BSMG

Possible to read just the correlation matrix into SAS for analysis.

```
data a (type = corr);
input _type_ $ _name_ $ gaelic english history math algebra geometry;
cards;
mean . 0 0 0 0 0 0
std . 1 1 1 1 1 1
n . 220 220 220 220 220 220
corr gaelic 1 .44 .41 .29 .33 .25
corr english .44 1 .35 .35 .32 .33
corr history .41 .35 1 .16 .19 .18
corr math .29 .35 .16 1 .59 .47
corr algebra .33 .32 .19 .59 1 .46
corr geometry .25 .33 .18 .47 .46 1 ;
run;
proc contents data = a; run;

proc factor data = a;
var gaelic english history math algebra geometry;
run;
```

Slide 18

options include method = prinit, method = uls, method = ml. By default with no method stated, SAS performs principal component analysis. Other options available are regarding the rotation method which can be used with all of the estimation methods. They are rotate=varimax and rotate=promax

School Subjects example from BSMG - PCA output

```
proc factor data = a rotate = promax;
var gaelic english history math algebra geometry; run;
```

The FACTOR Procedure

Initial Factor Method: Principal Components
 Prior Communality Estimates: ONE
 Eigenvalues of the Correlation Matrix: Total = 6 Average = 1

	Eigenvalue	Difference	Proportion	Cumulative
1	2.72868347	1.59989129	0.4548	0.4548
2	1.12879218	0.51350073	0.1881	0.6429
3	0.61529145	0.01248259	0.1025	0.7455
4	0.60280886	0.08029449	0.1005	0.8459
5	0.52251437	0.12060470	0.0871	0.9330
6	0.40190967		0.0670	1.0000

2 factors will be retained by the MINEIGEN criterion.
 Factor Pattern

	Factor1	Factor2
gaelic	0.66080	0.44447
english	0.68846	0.28977
history	0.51636	0.63955
math	0.73562	-0.41702
algebra	0.74187	-0.37276
geometry	0.67817	-0.35410

Variance Explained by Each Factor

	Factor1	Factor2
	2.7286835	1.1287922

Final Communality Estimates: Total = 3.857476

	gaelic	english	history	math	algebra	geometry
	0.63421803	0.55795080	0.67565080	0.71504052	0.68931666	0.58529884

Slide 19

School Subjects example from BSMG - PCA OUTPUT

Prerotation Method: Varimax

Orthogonal Transformation Matrix

	1	2
1	0.77371	0.63354
2	-0.63354	0.77371

Rotated Factor Pattern

	Factor1	Factor2
gaelic	0.22967	0.76254
english	0.34909	0.66037
history	-0.00567	0.82196
math	0.83335	0.14340
algebra	0.81015	0.18160
geometry	0.74904	0.15568

Variance Explained by Each Factor

	Factor1	Factor2
	2.0865260	1.7709497

Final Communality Estimates: Total = 3.857476

	gaelic	english	history	math	algebra	geometry
	0.63421803	0.55795080	0.67565080	0.71504052	0.68931666	0.58529884

Slide 20

School Subjects example from BSMG - PCA OUTPUT

Rotation Method: Promax (power = 3)

Inter-Factor Correlations

	Factor1	Factor2
Factor1	1.00000	0.36730
Factor2	0.36730	1.00000

Rotated Factor Pattern (Standardized Regression Coefficients)

	Factor1	Factor2
gaelic	0.09286	0.75757
english	0.23917	0.62522
history	-0.16759	0.86862
math	0.85268	-0.01983
algebra	0.82064	0.02525
geometry	0.76115	0.01045

Variance Explained by Each Factor Eliminating Other Factors

Factor1	Factor2
1.7940097	1.4883579

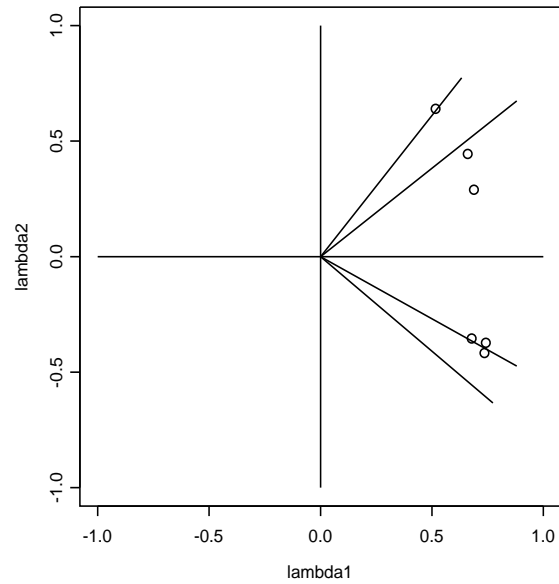
Variance Explained by Each Factor Ignoring Other Factors

Factor1	Factor2
2.3691177	2.0634659

Final Communality Estimates: Total = 3.857476

gaelic	english	history	math	algebra	geometry
0.63421803	0.55795080	0.67565080	0.71504052	0.68931666	0.58529884

Slide 21



Slide 22

School Subjects example from BSMG - ML OUTPUT

```
proc factor data = a method = ml rotate = promax;
var gaelic english history math algebra geometry; run;
```

The FACTOR Procedure

Initial Factor Method: Maximum Likelihood

Prior Communality Estimates: SMC

gaelic	english	history	math	algebra	geometry
0.30168715	0.29583987	0.20564352	0.41522406	0.41174029	0.29450749

Preliminary Eigenvalues: Total = 2.93847068 Average = 0.48974511

	Eigenvalue	Difference	Proportion	Cumulative
1	3.17047301	2.54848714	1.0790	1.0790
2	0.62198587	0.72927492	0.2117	1.2906
3	-.10728905	0.06189654	-0.0365	1.2541
4	-.16918559	0.08194532	-0.0576	1.1965
5	-.25113091	0.07525173	-0.0855	1.1111
6	-.32638264		-0.1111	1.0000

Significance Tests Based on 220 Observations

Test	DF	Chi-Square	Pr > ChiSq
H0: No common factors	15	308.5731	<.0001
HA: At least one common factor			
H0: 2 Factors are sufficient	4	2.1801	0.7027
HA: More factors are needed			

Slide 23

School Subjects example from BSMG - ML OUTPUT

Eigenvalues of the Weighted Reduced Correlation Matrix: Total = 5.71554705 Average = 0.95259117

	Eigenvalue	Difference	Proportion	Cumulative
1	4.57559833	3.43564960	0.8006	0.8006
2	1.13994873	1.05025464	0.1994	1.0000
3	0.08969409	0.06384279	0.0157	1.0157
4	0.02585130	0.03607174	0.0045	1.0202
5	-.01022044	0.09510453	-0.0018	1.0184
6	-.10532497		-0.0184	1.0000

Factor Pattern
Estimate/StdErr

	Factor1	Factor2
gaelic	0.55798 0.07342	0.42461 0.08368
english	0.56901 0.06148	0.28574 0.08579
history	0.39239 0.07933	0.44952 0.09364
math	0.73802 0.05948	-0.27943 0.06620
algebra	0.71847 0.04950	-0.20868 0.08511
geometry	0.59488 0.05347	-0.13331 0.08464

Slide 24

School Subjects example from BSMG - ML OUTPUT

Rotation Method: Promax (power = 3)

Rotated Factor Pattern (Standardized Regression Coefficients)

Estimate/StdErr

	Factor1	Factor2
gaelic	0.06268 0.07879	0.66737 0.09821
english	0.19141 0.08863	0.51810 0.09333
history	-0.08658 0.05532	0.63580 0.08912
math	0.81101 0.07027	-0.04511 0.05589
algebra	0.73469 0.07277	0.02597 0.06756
geometry	0.57413 0.07394	0.06505 0.07831

Slide 25

School Subjects example from BSMG - OUTPUT From MPLUS - ML

DATA:
FILE IS "C:\Mplus\ph5482\schoolsubjects.txt";
TYPE IS FULLCORR;
NOBSERVATIONS = 220;

VARIABLE:
NAMES ARE gaelic english history math algebra geometry;
USEVARIABLES ARE gaelic english history math algebra geometry;

ANALYSIS:
TYPE IS EFA 1 3;
ESTIMATOR IS ML;
ITERATIONS = 1000;
CONVERGENCE = 0.00005;

INPUT READING TERMINATED NORMALLY

SUMMARY OF ANALYSIS Number of groups
1 Number of observations
220

Number of dependent variables 6
Number of independent variables 0
Number of continuous latent variables 0
Observed dependent variables
Continuous
GAELIC ENGLISH HISTORY MATH ALGEBRA GEOMETRY

EIGENVALUES FOR SAMPLE CORRELATION MATRIX

	1	2	3	4	5	6
1	2.729	1.129	0.615	0.603	0.523	0.402

Slide 26

MPLUS - ML results for fitting 1 factor

RESULTS FOR EXPLORATORY FACTOR ANALYSIS

EXPLORATORY ANALYSIS WITH 1 FACTOR(S) :

CHI-SQUARE VALUE 52.682
 DEGREES OF FREEDOM 9
 PROBABILITY VALUE 0.0000

RMSEA (ROOT MEAN SQUARE ERROR OF APPROXIMATION) :
 ESTIMATE (90 PERCENT C.I.) IS 0.149 (0.111 0.189)
 PROBABILITY RMSEA LE 0.05 IS 0.000

ROOT MEAN SQUARE RESIDUAL IS 0.0987

ESTIMATED FACTOR LOADINGS

1

 GAELIC 0.500
 ENGLISH 0.539
 HISTORY 0.350
 MATH 0.726
 ALGEBRA 0.729
 GEOMETRY 0.615

ESTIMATED RESIDUAL VARIANCES

	GAELIC	ENGLISH	HISTORY	MATH	ALGEBRA	Geometry
1	0.750	0.710	0.878	0.473	0.468	0.621

Slide 27

MPLUS - ML results for fitting 2 factors

EXPLORATORY ANALYSIS WITH 2 FACTOR(S) :

CHI-SQUARE VALUE 2.233
 DEGREES OF FREEDOM 4
 PROBABILITY VALUE 0.6928

RMSEA (ROOT MEAN SQUARE ERROR OF APPROXIMATION) :
 ESTIMATE (90 PERCENT C.I.) IS 0.000 (0.000 0.077)
 PROBABILITY RMSEA LE 0.05 IS 0.854
 ROOT MEAN SQUARE RESIDUAL IS 0.0140

VARI MAX ROTATED LOADINGS

	1	2
GAELIC	0.660	0.236
ENGLISH	0.550	0.321
HISTORY	0.591	0.083
MATH	0.169	0.771
ALGEBRA	0.217	0.716
GEOMETRY	0.213	0.571

PROMAX ROTATED LOADINGS

	1	2
GAELIC	0.683	0.033
ENGLISH	0.533	0.167
HISTORY	0.648	-0.113
MATH	-0.031	0.805
ALGEBRA	0.040	0.726
GEOMETRY	0.077	0.566

PROMAX FACTOR CORRELATIONS

	1	2
1	1.000	
2	0.524	1.000

ESTIMATED RESIDUAL VARIANCES

	GAELIC	ENGLISH	HISTORY	MATH	ALGEBRA	Geometry
1	0.508	0.595	0.644	0.377	0.440	0.628

MUTHEN & MUTHEN 3463 Stoner Ave. Los Angeles, CA 90066 Tel: (310)

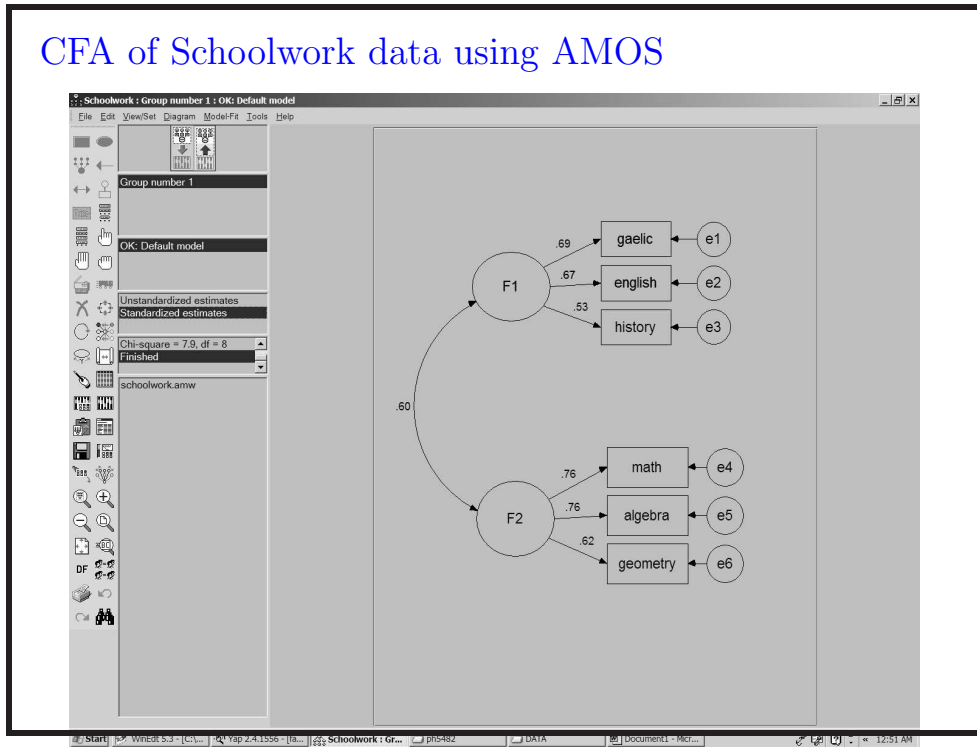
391-9971 Fax: (310) 391-8971 Web: www.StatModel.com Support:

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Slide 28

CFA of Schoolwork data using AMOS

Slide 29



CFA of Schoolwork data using MPLUS

Slide 30

```

DATA:
FILE IS "C:\Mplus\ph5482\schoosubjects.csv";
TYPE IS FULLCDD;
NOBSERVATIONS = 220;

VARIABLE:
NAMES ARE GAELIC ENGLISH HISTORY MATH ALGEBRA GEOMETRY;
USEVARIABLES ARE GAELIC ENGLISH HISTORY MATH ALGEBRA GEOMETRY;

ANALYSIS:
TYPE IS GENERAL;
ESTIMATOR IS ML;
ITERATIONS = 1000;
CONVERGENCE = 0.00005;

model:
f1 by GAELIC ENGLISH HISTORY;
f2 by MATH ALGEBRA GEOMETRY;

OUTPUT: STANDARDIZED;
-----
Number of observations                220
Number of dependent variables         6
Number of independent variables       0
Number of continuous latent variables 2

Observed dependent variables
Continuous
  GAELIC   ENGLISH   HISTORY   MATH   ALGEBRA   GEOMETRY

Continuous latent variables
  F1       F2

Estimator                                ML
    
```

CFA of Schoolwork data using MPLUS

```

TESTS OF MODEL FIT

Chi-Square Test of Model Fit
  Value          7.954
Degrees of Freedom      8
P-Value          0.4379

Chi-Square Test of Model Fit for the Baseline Model
  Value          314.045
Degrees of Freedom     15
P-Value          0.0000

CFI/TLI
  CFI          1.000
  TLI          1.000

Loglikelihood
  H0 Value     -1716.947
  H1 Value     -1712.969

Information Criteria
  Number of Free Parameters      13
  Akaike (AIC)                  3459.893
  Bayesian (BIC)                3504.010
  Sample-Size Adjusted BIC      3462.813
  (n* = (n + 2) / 24)

RMSEA (Root Mean Square Error Of Approximation)
  Estimate          0.000
  90 Percent C.I.    0.000  0.079
  Probability RMSEA <= .05    0.764

SRMR (Standardized Root Mean Square Residual)
  Value            0.033
    
```

Slide 31

CFA of Schoolwork data using MPLUS

```

              Estimates   S.E.  Est./S.E.  Std  StdYX

F1      BY
  GAELIC      1.000   0.000   0.000   0.688   0.690
  ENGLISH     0.973   0.151   6.434   0.670   0.671
  HISTORY     0.770   0.133   5.806   0.530   0.531

F2      BY
  MATH        1.000   0.000   0.000   0.762   0.764
  ALGEBRA     1.000   0.116   8.630   0.762   0.764
  GEOMETRY    0.806   0.104   7.757   0.615   0.616

F2      WITH
  F1          0.314   0.062   5.075   0.598   0.598

Variances
  F1          0.474   0.104   4.568   1.000   1.000
  F2          0.581   0.103   5.667   1.000   1.000

Residual Variances
  GAELIC      0.522   0.082   6.381   0.522   0.524
  ENGLISH     0.547   0.081   6.752   0.547   0.550
  HISTORY     0.715   0.081   8.772   0.715   0.718
  MATH        0.414   0.068   6.085   0.414   0.416
  ALGEBRA     0.415   0.068   6.091   0.415   0.416
  GEOMETRY    0.618   0.071   8.681   0.618   0.620

R-SQUARE
  Obs.Variable  R-Square
  GAELIC        0.476
  ENGLISH       0.450
  HISTORY       0.282
  MATH          0.584
  ALGEBRA       0.584
  GEOMETRY      0.380
    
```

Slide 32

The chi-square test for factor analysis models - EFA

- The chi-square test can also be used to see how well an EFA model with q factors fits the data
- For the EFA model the d.f. are specifically $\frac{p(p+1)}{2} - pq - p + \frac{q(q-1)}{2}$
- In EFA, the model being tested is
 - H_0 : The q factor model is correct (i.e. q factors sufficiently describe the p dimensional vector)
 - H_A : More factors are needed
- From Bartholomew and Knott (1998) “Starting with $q = 1$, we then take successive values in turn until the fit of the model is judged to be adequate. Viewed as a testing procedure this is not strictly valid because it does not adjust the significance levels to allow for the sequential character of the test. It rather depends on regarding the p-value of the test as a measure of the adequacy of the model”
- Because of the tendency for the approach described just above to keep adding more and more factors (since more factors will make the model fit better), BK suggest considering the AIC Akaike’s information criterion. This is simply a penalized version of the log likelihood where models with more parameters are penalized.

Slide 33

The chi-square test for factor analysis models - CFA

- For CFA the d.f. are $p(p+1)/2$ – (the number of parameters being estimated)
- In CFA, the model being tested is
 - H_0 : The model being assumed is correct
 - H_A : A more complicated model is needed
- Satorra-Bentler adjustment, Bollen-Stine adjustment, Amemiya Anderson result

Slide 34

Things we’ll do next - Read all of BSMG Chapter 6 and Kline Chapter 7

- More about CFA
- equivalent models
- Nested models
- Multiple Group