

# On the Behavior of Marginal and Conditional Akaike Information Criteria in Linear Mixed Models

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# Overview

- Additive mixed models offer growing flexibility in modeling, making model choice increasingly important.
- Two Akaike Information Criteria (AICs) have been used in mixed models: the marginal and the conditional AIC.
- We shed light on their theoretical and practical properties when [selecting random effects](#).

1 Background: Mixed Models and the AIC

2 Marginal and Conditional AIC

3 Application: Childhood Malnutrition in Zambia

# Linear and Additive Mixed Models

Linear Mixed Models (LMMs): widely used class of flexible regression models

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{b} + \boldsymbol{\varepsilon}, \quad \begin{pmatrix} \mathbf{b} \\ \boldsymbol{\varepsilon} \end{pmatrix} \sim \mathcal{N} \left( \begin{pmatrix} \mathbf{0} \\ \mathbf{0} \end{pmatrix}, \begin{pmatrix} \mathbf{D} & \mathbf{0} \\ \mathbf{0} & \sigma^2 \mathbf{I}_n \end{pmatrix} \right).$$

**Special case I:** Random intercept / random slope models for [longitudinal](#) or [grouped data](#) → model intra-subject correlations.

**Special case II:** [Penalized spline smoothing](#) using a particular LMM:

- polynomial part (degree  $d$ ) of smooth function corresponds to fixed effects
- deviations modeled using random effects.

Similar approaches for varying coefficients, surfaces, spatial effects, ...

→ [Additive mixed models](#) can combine all of these modeling components.

# Application: Childhood Malnutrition in Zambia

- 1,600 observations from the 1992 Zambia Demographic and Health Survey
- Additive mixed model for stunting, measuring insufficient height for age:

$$zscore_i = \mathbf{x}'_i \boldsymbol{\beta} + m_1(cage_i) + m_2(cfeed_i) + m_3(mage_i) + m_4(mbmi_i) \\ + m_5(mheight_i) + b_{s_i} + \varepsilon_i,$$

- Categorical covariates: gender of the child, education and employment status of the mother
- Continuous covariates: age of the child, duration of breastfeeding, age, height and body mass index of the mother - **smooth or linear effects?**
- Random intercept:  $b_{s_i}$  captures **spatial variability** between residential districts.
- Model **selection** involves **random effects**.
  - Random intercept
  - Random effects modeling smooth deviations from linearity of the  $m_k(\cdot)$ .

# The Akaike Information Criterion (AIC)

- Data  $\mathbf{y}$  generated from a true underlying model with density  $g(\cdot)$ .  
Approximate by a parametric class of models  $f_{\psi}(\cdot) = f(\cdot; \psi)$ ,  $\psi \in \Psi$ .
- Kullback-Leibler (KL) distance:  $K(f_{\psi}, g) = E_{\mathbf{z}} [\log(g(\mathbf{z})) - \log(f_{\psi}(\mathbf{z}))]$ .
- Want to minimize (relative) expected KL distance using estimator  $\hat{\psi}(\mathbf{y})$

$$E_{\mathbf{y}}[K(f_{\hat{\psi}(\mathbf{y})}, g)] \rightarrow \min \quad \Leftrightarrow \quad -2 E_{\mathbf{y}}[E_{\mathbf{z}} [\log(f_{\hat{\psi}(\mathbf{y})}(\mathbf{z}))]] \rightarrow \min \quad (1)$$

Predictive quantity, with  $\mathbf{y}$  and  $\mathbf{z}$  independent replications.

- $AIC = -2 \log(f_{\hat{\psi}(\mathbf{y})}(\mathbf{y})) + 2k$  is an estimator of (1),  $k = \dim(\psi)$ .
  - $2k$  is correction term:  $-2 \log(f_{\hat{\psi}(\mathbf{y})}(\mathbf{y}))$  not predictive quantity.
  - unbiased for (1) under regularity conditions:  $\Psi = \mathbb{R}^k$  and  $y_1, \dots, y_n$  are i.i.d.

... and for the linear mixed model?

# Marginal and Conditional Perspective

- **Marginal perspective:** Random effects induce a **correlation structure**.

$$\mathbf{y} \sim \mathcal{N}(\mathbf{X}\boldsymbol{\beta}, \mathbf{V}), \quad \mathbf{V} \equiv \mathbf{V}(\boldsymbol{\theta}) = \sigma^2 \mathbf{I}_n + \mathbf{ZDZ}'.$$

$$mAIC = -2l(\mathbf{y}|\hat{\boldsymbol{\beta}}, \hat{\boldsymbol{\theta}}) + 2(p + q), \quad \text{where}$$

$l(\mathbf{y}|\hat{\boldsymbol{\beta}}, \hat{\boldsymbol{\theta}})$  is the maximized marginal log-likelihood,  $p = \dim(\boldsymbol{\beta})$ ,  $q = \dim(\boldsymbol{\theta})$ .

Typically returned by statistical **software** (R, SAS) and widely used for selection of random and fixed effects.

- **Conditional perspective:** Random effects are **regression coefficients** estimated subject to regularization.

$$\mathbf{y}|\mathbf{b} \sim \mathcal{N}(\mathbf{X}\boldsymbol{\beta} + \mathbf{Zb}, \sigma^2 \mathbf{I}_n).$$

# The Conditional AIC

- Vaida & Blanchard (Biometrika, 2005): **Conditional AIC**

$$cAIC = -2l(\mathbf{y}|\hat{\boldsymbol{\beta}}, \hat{\mathbf{b}}, \hat{\boldsymbol{\theta}}) + 2(\rho + 1), \quad \text{where}$$

$l(\mathbf{y}|\boldsymbol{\beta}, \mathbf{b}, \boldsymbol{\theta})$  is the conditional log-likelihood and  $\rho$  is the trace of the hat matrix, the **effective degrees of freedom**, when  **$D$  known**.

Authors recommend the cAIC with estimated  $\hat{D}$  when  $D$  is unknown:  $\hat{\rho}$ .

- Liang, Hu & Zou (Biometrika, 2008): **Corrected cAIC** replaces  $\rho$  by  $\Phi_0 = \text{trace} \left( \frac{\partial \hat{\mathbf{y}}}{\partial \mathbf{y}} \right)$  for known  $\sigma^2$ . (Otherwise,  $\Phi_1$  involves second derivatives.)

No closed form, numerical approximation of  $\Phi_0$  requires  **$n$  extra model fits**.  
In our example: ca. 110 days (!) to compute corrected cAIC for all models.

# Questions

- Theoretical and practical properties of the marginal and conditional AIC when focus is on the **random effects**?
- Can we use the computationally less demanding cAIC of Vaida & Blanchard?

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# Theoretical Results

## Marginal AIC

- The marginal AIC is positively biased for (1).
- The bias is dependent on the true unknown  $\theta$ , does not vanish asymptotically.
- The marginal AIC favors smaller models excluding random effects.

**Reasons:**  $\Psi \neq \mathbb{R}^{p+q}$  and  $y_1, \dots, y_n$  correlated  $\Rightarrow$  bias correction not  $2 E[\chi_{\rho+q}^2]$ .

Compare [testing for zero variances](#) (Crainiceanu & Ruppert, 2004; Greven et al, 2008).

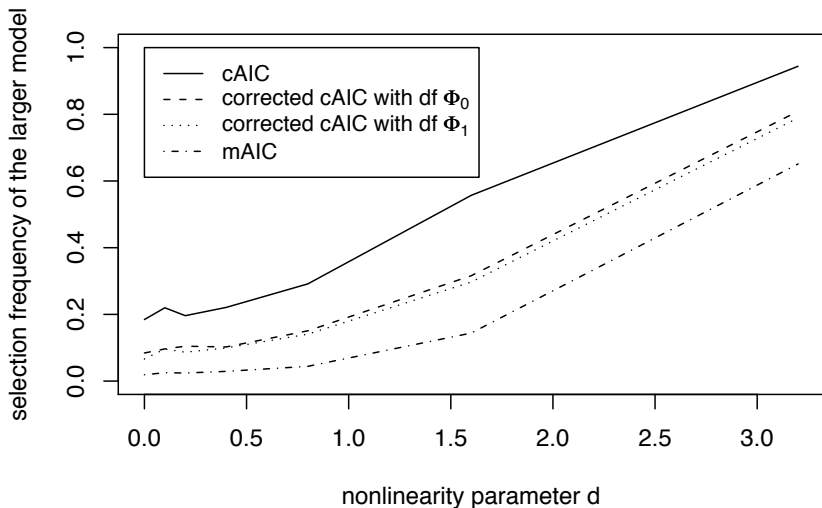
## Conditional AIC of Vaida & Blanchard with estimated $\rho$

- The larger model including the random effect is chosen *whenever*  $\hat{\mathbf{b}} \neq 0$ .
- If  $\hat{\mathbf{b}} = 0$ , the cAICs of the two models coincide.

**Reasons:** Bias correction  $\rho$  again estimated from the same data  $\mathbf{y}$ .

## Simulations - Example

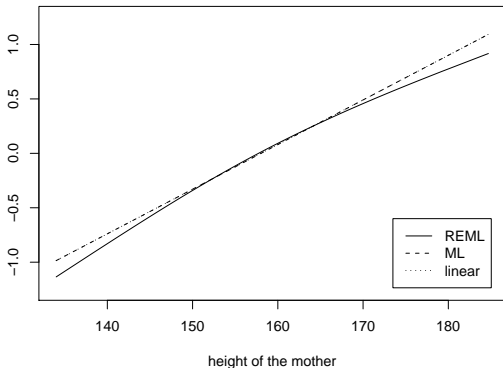
$y_i = m(x_i) + \varepsilon_i, i = 1, \dots, n = 30$ , where  $m(x) = 1 + x + 2d(0.3 - x)^2$ .



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# Application: Childhood Malnutrition in Zambia

Simple example:  $zscore_i = m(mheight_i) + \varepsilon_i$ .



ML: mAIC prefers smaller, linear model; tie for cAIC.

REML: mAIC prefers smaller, linear model; cAIC chooses larger, non-linear model.

In the full model the functions are either linear or clearly non-linear . . . .

# Summary

- The **marginal AIC** is affected by the boundary of the parameter space, similarly to likelihood ratio tests.  
It is biased towards **simpler models** excluding random effects.
- The **conventional conditional AIC** tends to select models that are **too large**.  
It includes any random effect not estimated to be exactly zero.
- The **corrected conditional AIC** **rectifies** this difficulty but comes at a high computational price.  
→ Work in progress: Representation that is computationally feasible?

## References

- Greven, S. & Kneib, T. (2009): On the Behavior of Marginal and Conditional Akaike Information Criteria in Linear Mixed Models. *Johns Hopkins University, Department of Biostatistics Working Papers, Paper 179*.  
<http://www.bepress.com/jhubiostat/paper179/>
- Akaike, H. (1973): Information theory and an extension of the maximum likelihood principle. In B. N. Petrov and F. Csaki (Eds.), *2nd International Symposium on Information Theory*, 267-281. Akademiai Kiado.
- Crainiceanu, C. & Ruppert, D. (2004): Likelihood ratio tests in linear mixed models with one variance component. *Journal of the Royal Statistical Society: Series B* 66(1), 165-185.
- Greven, S., Crainiceanu, C., Küchenhoff, H. & Peters, A. (2008): Restricted likelihood ratio testing for zero variance components in linear mixed models. *Journal of Computational and Graphical Statistics* 17 (4): 870-891.
- Liang, H., Wu, H. & Zou, G. (2008): A note on conditional AIC for linear mixed-effects models. *Biometrika* 95, 773-778.
- Vaida, F. & Blanchard, S. (2005): Conditional Akaike information for mixed-effects models. *Biometrika* 92, 351-370.