

Posterior consistency for some Bayesian Semi-parametric Problems

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1 Introduction

2 Posterior Consistency for some Semi-parametric Problems

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 - 1 More flexible than parametric models.
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 - 1 Reflecting one's prior beliefs into the analysis.
 - 2 Straightforward in principle where inference is based on the posterior distribution only.

Priors on Infinite-dimensional Spaces

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 - Each realization of P is a probability measure on $(\mathcal{X}, \mathcal{B})$;
 - For each measurable finite partition $\{B_1, \dots, B_k\}$ of \mathcal{X} , the joint distribution of the vector $(P(B_1), \dots, P(B_k))$ has Dirichlet distribution on the k -dimensional uni simplex with parameters $(k; \alpha(B_1), \dots, \alpha(B_k))$.

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 - $P \sim \mathcal{D}_\alpha$: Random mixing distribution on Θ ;
 - \mathcal{D}_α : a prior on $\mathcal{M}(\Theta)$, the space of probability measures on Θ
 - Via the map $P \mapsto f_P(x) := \int K(x; \theta) dP(\theta)$, there is a prior on $\mathcal{D}(\mathcal{X})$, the space of densities on \mathcal{X} .

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- Gaussian process Lévy process (Leonard (1978), Lenk (1988,1991), Ferguson and Phadia (1979) and Lo (1982)).

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Definition

For each n , let $\Pi(\cdot|\mathbb{X}_n)$ be a posterior given X_1, \dots, X_n . The sequence $\{\Pi(\cdot|\mathbb{X}_n)\}$ is said to be consistent at θ_0 if there is a $\Omega_0 \subset \Omega$ with $P_{\theta_0}^\infty(\Omega_0) = 1$ such that if $\omega \in \Omega_0$, then for every neighborhood U of θ_0 ,

$$\Pi(U|\mathbb{X}_n(\omega)) \rightarrow 1. \quad (1)$$

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- Doob (1948): The prior probability of the set of all $\theta \in \Theta$ at which consistency does not hold, is 0. This result does not imply consistency at a given point if there is no prior point mass there.
- Freedman (1963, 1986): Only having positive prior probability in a neighborhood of the true value of the parameter does not imply consistency at that value.

Basic Theorem about Posterior Consistency

■ Schwartz (1965)

Theorem

Let $\theta_0 \in U \subset \Theta$. If there exists $m \geq 1$, a test function $\phi(\mathbb{X}_m)$ for testing $H_0 : \theta = \theta_0$ against $H_a : \theta \in U^c$ with the property that $\inf\{E_\theta \phi(\mathbb{X}_m) : \theta \in U^c\} > E_{\theta_0} \phi(\mathbb{X}_m)$ and $\theta_0 \in KL(\Pi)$, then $\Pi\{\theta \in U^c | \mathbb{X}_n\} \rightarrow 0$ a.s. $[P_{\theta_0}^\infty]$.

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- Test condition
- KL property

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$$\mathcal{H}_\epsilon(f) = \{g : \mathcal{K}(f, g) < \epsilon\}.$$
- We say that the KL property holds at f_0 or f_0 is in the Kullback-Leibler support of Π^* and write $f_0 \in KL(\Pi^*)$, if $\Pi^*(\mathcal{H}_\epsilon(f_0)) > 0$ for every $\epsilon > 0$.

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- My work:
Studying posterior consistency for various semi-parametric models.

General Model Notations

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- $a_0(\cdot)$ and θ_0 : the true values of the parameters. Let $f_0 = f_{a_0,\theta_0}$ and $f_{0i} = f_{a_0,\theta_0,i}$.

Main Theorem

Theorem

Let $\mathcal{W} \subset \mathcal{A} \times \Theta$ and let X_i 's be i.i.d. with distribution function Q .
If for some sequences $\mathcal{A}_n \subset \mathcal{A}$ and $m_n \leq M_n$,

- (i) there is an exponentially consistent sequence of tests for testing: $(a_0(\cdot), \theta_0)$ v.s. $\mathcal{W} \cap \left(\mathcal{A}_n \times [m_n, M_n]^d \right)$;
- (ii) Kullback Leibler property holds;
- (iii) the prior measure of the outside of the sieve is exponentially small;

then with $(P_{f_0}^\infty)$ -probability 1, the model is consist on $(a_0(\cdot), \theta_0)$,
i.e.

$$\Pi(\mathcal{W} | (X_1, Y_1), \dots, (X_n, Y_n)) \rightarrow 0. \quad (2)$$

Exponential consistency:

For i.i.d. X_i 's, there exist constants $C_1, C_2, C > 0$, such that

$$\int [E_{f_0^n}(\Phi_n)] Q^n(dx) \leq C_1 e^{-nC}, \text{ and}$$

$$\inf_{(a(\cdot), \theta) \in \mathcal{W}} \int [E_{f_{a, \theta}^n}(\Phi_n)] Q^n(dx) \geq 1 - C_2 e^{-nC}.$$

Consistency of Multiple Regression

Generalization of Theorem 4.1 in Amewou-Atiss *et al.* (2003).

Theorem

Suppose that

- (i) the covariate X is compactly supported, and for any quadrant Γ , $Q(\Gamma \setminus \{|X_i| < \zeta\}) > 0$ for each $i = 1, \dots, d$, and some $\zeta > 0$;
- (ii) f_0 is continuous, $f_0(0) > 0$, and f_0 is covered by some other function: for $|\eta'| < \eta$, $f_0(y - \eta') < C_\eta g_\eta(y)$ for all y ;
- (iii) $\tilde{\Pi}\{K(g_\eta, f) < \xi\} > 0$ and (α_0, β_0) is in the support of μ .

Then for any weak neighborhood \mathcal{U} of f_0 , $\Pi\left\{f \in \mathcal{U}, |\alpha - \alpha_0| < \epsilon, \|\beta - \beta_0\| < \epsilon \mid (X_1, Y_1), \dots, (X_n, Y_n)\right\} \rightarrow 1$ a.s. $P_{f_0}^\infty$.



Exponential Frailty Model

- Let $(X_1, Y_1), (X_2, Y_2), \dots$ be paired observations.

$$\begin{aligned} X|W = w &\sim \text{Exp}(w), \quad Y|W = w \sim \text{Exp}(\lambda w), \\ \lambda &\sim \mu, \quad W \sim F, \quad F \sim \tilde{\Pi}, \end{aligned}$$

where $\text{Exp}(w)$ is corresponding to p.d.f. we^{-wx} .

For given $W = w$, X and Y are conditionally independent, $\tilde{\Pi}$ and μ are independent.

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- The likelihood function of observations:

$$f(x, y) = \int \lambda w^2 e^{-w(x+\lambda y)} dF(w)$$

Consistency of Exponential Frailty Model

Theorem

Suppose that $\log(f_0(x, y))$, x and y are f_0 -integrable, F_0 is in the weak support of $\tilde{\Pi}$, λ_0 is in the support of μ and

$w_E := \int w^2 dF_0(w) < \infty$. Let

$\mathscr{W} = \{(F, \lambda) : |\lambda - \lambda_0| < \epsilon, F \in \mathscr{U}\}^c$, where \mathscr{U} is a weak neighborhood of F_0 . If for any $\delta > 0$ and $\beta > 0$ there exists a

sequence r_n and a constant β_0 such that $r_n^2 < n\beta$ and

$\tilde{\Pi}\{F : F(r_n) - F(r_n^{-1}) < 1 - \delta\} < e^{-n\beta_0}$, then

$\Pi\{\mathscr{W} | (X_1, Y_1), (X_2, Y_2), \dots, (X_n, Y_n)\} \rightarrow 0$ a.s. $\prod_{i=1}^{\infty} P_{f_{0i}}$.

Other studies about this model: Cantor *et al.* (1985) and Owen *et al.* (2000).

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- $f(t) = h(t) \exp(-H(t)) = h(t) \bar{F}(t)$: baseline random density function for life time Y .

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- $f(t) = h(t) \exp(-H(t)) = h(t) \bar{F}(t)$: baseline random density function for life time Y .
- Censoring: (Z, Δ) , where $Z = Y \wedge C$, $\Delta = \mathbb{1}(Y \leq C)$ for C a censoring time with distribution F_C .

Consistency of Cox Proportional Hazard Model

Theorem

If

- (a1) $f_0(t)$ is strictly positive on $(0, \infty)$ and $\int_{\mathbb{R}^+} \max\{E[H(t)^q], t^q\} f_0(t) dt < \infty$ for any $q > 1$;
- (a2) there exists $r > 0$ such that $\liminf_{t \downarrow 0} h(t)/t^r = \infty$, a.s.;
- (a3) $\{h : \sup_{0 < t < T} |h(t) - h_0(t)| < \delta, \int_T^\infty |H(t) - H_0(t)| f_0(t) dt < \delta\}$ has positive prior probability for some finite T and $\delta > 0$;
(Blasi et al. (2009))
- (a4) $\int_0^\infty \log(f_0(t))^q f_0(t) dt < \infty$.

Then, for $\mathscr{W} = \{(h, \beta) : |\beta - \beta_0| < \epsilon, h : 1 - \exp(-H(t)) \in \mathcal{U}_{F_0}\}^c$, we have that $\Pi(\mathscr{W} | (X_1, Y_1), \dots, (X_n, Y_n)) \rightarrow 0$ a.s. $[P_{\beta_0, h_0}^\infty]$.

Generalized Linear Models with Unknown Link Functions

Traditionally, in a generalized linear model (GLM), a known function of the expectation of the observations, called the link function, is modeled to be linear in the predictors. Recently, for more flexible modeling, the link function is treated as unknown or the linearity assumption is removed. In the Bayesian context, the former is studied by Gelfand and Kuo (1991), Newton *et al.* (1996), Mallick and Gelfand (1994), Basu and Mukhopadhyay (2000) and others.

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- $F(X_i^T \beta) := T(g(X_i^T \beta))$;
- Prior: $F \sim \tilde{\Pi}$, $\beta \sim \mu$;
- $\|\beta\| = 1$.

Consistency of Generalized Linear Models with Unknown Link Functions

Theorem

The covariate X is compactly supported, and for any quadrant Γ , $Q(\Gamma \setminus \{X : |X_i| < \zeta\}) > 0$ for each $i = 1, \dots, d$, and some $\zeta > 0$; Assume that the weak support of $\tilde{\Pi}$ covers $\mathcal{M}([-L, L])$. If $\pi(B_n^c) \leq c_1 e^{-nc_2}$ for some constants $c_1, c_2 > 0$, then with $P_{f_0}^\infty$ -probability 1, the posterior probability

$$\Pi(\mathcal{W} | (X_1, Y_1), \dots, (X_n, Y_n)) \rightarrow 0,$$

where \mathcal{W} is of the form $(\mathcal{U}_{F_0} \times \{\beta : \|\beta - \beta_0\| < \delta\})^c$ and \mathcal{U}_{F_0} is the weak neighborhood of F_0 .

Accelerated Failure Time Model

- For each subject $i = 1, \dots, n$, let Y_i denote the failure time and C_i denote the censoring time. The observed survival data are $Z_i = \min(Y_i, C_i)$ and $\Delta_i = \mathbb{1}(Y_i \leq C_i)$. Let X_i denote a d -dimensional vector of covariates associated with subject i .

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- $\beta \sim \mu$.
- $\tilde{\Pi} \times \nu \times \mu$ induces a prior on the density function \tilde{f} of $\log Y - X^T \beta$ via the map

$$(P, \theta) \mapsto \int \theta \exp(-e^{(t - \log \phi)\theta} + (t - \log \phi)\theta) dP.$$

Sufficient Conditions for Consistency of AFT

1. the covariate X is compactly supported, and for any quadrant Γ , $Q(\Gamma \setminus \{\|x_i\| < \zeta\}) > 0$, for all $i = 1, \dots, d$ and some $\zeta > 0$;
2. for some $0 < M < \infty$, $0 < \tilde{f}_0(t) < M$ for all t ;
3. $\int \tilde{f}_0(t) |\log \tilde{f}_0(t)| dt < \infty$ and $\int \tilde{f}_c(t + \xi) |\log \tilde{F}_0(t)| dt < \infty$ for all $\xi \in \mathbb{R}$;
4. for some $\delta > 0$, $\int \tilde{f}_0(t) \log \frac{\tilde{f}_0(t)}{\phi_\delta(t)} dt < \infty$;
5. there exists $\eta > 0$ such that $|\int e^{2t|t|^\eta} \tilde{f}_0(t) dt| < \infty$, and $\int \tilde{f}_0(t) e^{(t-a)/b} dt < \infty$ for all $a \in \mathbb{R}$ and $b \in (0, \infty)$;
6. $\int (e^{-t/h} + t/h) dP(t) < \infty$ for any given $h > 0$, a.s. $\tilde{\Pi}$;
7. the weak support of $\tilde{\Pi}$ is the space of all probability measures on \mathbb{R} ;
8. for some $\delta > 0$ and any $\xi \in \mathbb{R}$, $\int \tilde{f}_c(t + \xi) \log \frac{\tilde{F}_{0, \beta_0}(t)}{\phi_\delta(t)} dt < \infty$;
9. there exists $\eta > 0$ such that $|\int \tilde{f}_c(t + \xi) e^{2t|t|^\eta} dt| < \infty$ for any $\xi \in \mathbb{R}$, and $\int \tilde{f}_c(t) e^{(t-a)/b} dt < \infty$ for all $a \in \mathbb{R}$ and $b \in (0, \infty)$;
10. $\int e^{t/h} \tilde{f}_0(t) dt < \infty$ and $\int e^{t/h} \tilde{f}_c(t) dt < \infty$ for any given $h > 0$.

Partial Linear Regression

- Model: $Y_i = X_i^T \beta + f(Z_i) + \epsilon_i$,
- ϵ_i 's are i.i.d. $\text{Norm}(0, \sigma^2)$,
- f is an odd function, random.
- Prior: $\beta \sim \mu$ $q_{f,z} * \phi(0, \sigma^2) \sim \tilde{\Pi}$, $\Pi = \tilde{\Pi} \times \mu$.

Theorem

Suppose that

- (1) X and Z are compactly supported, and distributed well;
- (2) Cover: $q_{f_0,z} * \phi(0, \sigma_0^2)(y - \eta') < C_\eta g_\eta(y)$;
- (3) KL property: $\tilde{\Pi}\{K(g_\eta, f) < \delta\} > 0$.

Then $\Pi\{(f, \beta) : f \in \mathcal{U}, \|\beta - \beta_0\| < \delta | (X_1, Y_1, Z_1), \dots, (X_n, Y_n, Z_n)\} \rightarrow 1$
a.s. $[P_{f_0}^\infty]$, for any weak neighborhood \mathcal{U} of f_0 .

The End