

# Chapter 9 Refinements of Semi-parametric PHM

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## §9.2 Time-dependent covariates

- $Z$ : baseline/fixed, not changing over time; most common.  
 $Z(t)$ : time-dependent or time-varying covariate, e.g. weight, BMI, blood pressure, health or trt status at time  $t$ .
- PHM and PL:

$$h(t|Z(t)) = h_0(t) \exp(Z(t)' \beta).$$

$$L(\beta) = \prod_{i=1}^D L_i = \prod_{i=1}^D \frac{\exp(Z_{(i)}(t)' \beta)}{\sum_{j \in R(t_i)} \exp(Z_{(j)}(t)' \beta)}.$$

- Example:  $n = 4$ ;  $Z(t)$ : smoking status at  $t$ .  
obs 1:  $Z(t) = 1$  for  $0 \leq t < 2$ ;  $Z(t) = 0$  for  $t \geq 2$ ; event at  $t = 5$ .  
obs 2:  $Z(t) = 0$  for  $0 \leq t < 4$ ;  $Z(t) = 1$  for  $t \geq 4$ ; event at  $t = 6$ .  
obs 3:  $Z(t) = 0$  up to an event at  $t = 1$ .  
obs 4:  $Z(t) = 1$  up to an event at  $t = 3$ .

$$L_1 = \frac{\exp(0 * \beta)}{\exp(0 * \beta) + \exp(1 * \beta) + \exp(0 * \beta) + \exp(1 * \beta)}.$$

$$L_2 = \frac{\exp(1*\beta)}{\exp(1*\beta) + \exp(0*\beta) + \exp(1*\beta)}.$$

$$L_3 = \frac{\exp(0*\beta)}{\exp(0*\beta) + \exp(1*\beta)}.$$

$$L_4 = \frac{\exp(1*\beta)}{\exp(1*\beta)} = 1.$$

- Example 9.1: SAS

DFS: disease-free survival, min(disease recurrence, death).

Dan Sargent's paper: in colon-cancer trials, the outcome is

- 1) 5-yr survival as the current practice;
- 2) 3-yr DFS as the new proposed one.

$Z_a = 0$  if  $t <$  time at which aGVHD occurs;  $= 1$  o/w.

$Z_c = 0$  if  $t <$  time at which cGVHD occurs;  $= 1$  o/w.

$Z_p = 0$  if  $t <$  time at the platelets recovered;  $= 1$  o/w.

- An application: check the PH assumption.
- Consider a simple example: a binary  $Z = 0$  or  $1$ , indicating one of the two groups.

PHM:  $h(t|Z) = h_0(t) \exp(Z\beta_1) \implies$

$$\frac{h(t|\text{Grp 1})}{h(t|\text{Grp 0})} = \exp(\beta_1), \text{ constant!}$$

- Now, consider an alternative model:

$$h(t|Z) = h_0(t) \exp(Z\beta_1 + Zg(t)\beta_2) \implies$$

$$\frac{h(t|\text{Grp 1})}{h(t|\text{Grp 0})} = \exp(\beta_1 + g(t)\beta_2), \text{ time-varying (if } \beta_2 \neq 0\text{)!}$$

- Checking a PHM: add a time-dependent covariate, e.g.  $Z \times \log(t)$ , and see whether its coefficient is significant!  
not a general GOF test: depending on the true and specified  $g(t)$ ; you used this type of tests often...
- Example 9.2: SAS

## §9.3 Stratified PHM

- Consider two binary covariates:  $Z_1$  indicating trt grp,  $Z_2$  indicating male grp.

$$\text{PHM: } h(t|Z_1, Z_2) = h_0(t) \exp(Z_1\beta_1 + Z_2\beta_2) \implies$$

$$\frac{h(t|trt, Z_2)}{h(t|ctl, Z_2)} = \exp(\beta_1),$$

$$\frac{h(t|Z_1, M)}{h(t|Z_1, F)} = \exp(\beta_2). \text{ --again constant and PH!}$$

- How about if 1) PH assumption may not hold for  $Z_2$  while holds for  $Z_1$  and 2) only  $\beta_1$  is of interest?

- 1) fix the problem, e.g., by ...; 2) use a *stratified PHM*:

$$h(t|Z_1, Z_2) = h_0(t|Z_2) \exp(Z_1\beta_1)$$

- Implications:

$$h(t|Z_1, M) = h_{0,M}(t) \exp(Z_1\beta_1),$$

$$h(t|Z_1, F) = h_{0,F}(t) \exp(Z_1\beta_1),$$

and there is **no** restriction/assumption on the relationship between the two baseline hazards; e.g.,

$h(t|Z_1, M)/h(t|Z_1, F) = h_{0,M}(t)/h_{0,F}(t)$  may or may *not* be a constant!

On the other hand,

$\frac{h(t|trt, Z_2)}{h(t|ctl, Z_2)} = \exp(\beta_1)$ , still a PHM!

- Summary: *to adjust for a confounder*, 1) put it in a regression model; or 2) treat it as a stratifier.

A key difference: 1) may require a stronger modeling assumption.

- How to estimate  $\beta$ ?
- Method 1: analogous to M-H method,
  - 1) fit a PHM for each stratum: using  $L^{(h)}(\beta)$  based on the data in stratum  $h \implies \hat{\beta}^{(h)}$ ;
  - 2) combine  $\hat{\beta}^{(h)}$ 's, e.g., by ...

Prior to 2), may conduct a homogeneity test on  $H_0$ :

$\beta^{(1)} = \beta^{(2)} = \dots$ ; why and how?

- Method 2: under the homogeneity assumption, use  $L(\beta) = \prod_h L^{(h)}(\beta)$ .
- Example 9.1b: SAS

## §9.4 Left-truncated and right-censored data

- Channing House data:  
Age 1: age at the entry;  
Age 2: age at the end of study or death.  
Covariate: gender.
- Recall how we dealt with LT-RC'ed data:  
One sample problem: modify K-M or N-A estimator; how?  
K-sample problem: modify generalized rank tests; how?  
Regression with PHM: modify PL; how?
- LT-RT'ed data:  $(Y_j, T_j, \delta_j, Z_j), j = 1, \dots, n.$
- PHM:  $h(t|Z) = h_0(t) \exp(Z'\beta).$
- Modification:  $R(t_i) = \{j : Y_j < t_i \leq T_j\};$   
PL remains the same with the new  $R(t_i).$
- MPLE



$$\hat{\beta} = \operatorname{argmax}_{\beta} \prod_{i=1}^D \frac{\exp(Z'_{(i)}\beta)}{\sum_{j \in R(t_i)} \exp(Z'_j\beta)},$$

$$\hat{H}_0(t) = \sum_{t_i \leq t} \frac{d_i}{\sum_{j \in R(t_i)} \exp(Z'_j\beta)}.$$

- Example 9.4: SAS
- Another use of the counting process notation:  
Code time-dependent covariates: e.g., how to input the data with time-dependent smoking status in the example on p.2?  
Any issue with multiple "observations" from the same subject?