

A novel and efficient algorithm for de novo discovery of mutated driver pathways in cancer

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Introduction

- ▶ Given: an $n \times p$ mutation matrix \mathbf{A} with entry $A_{ij} = 1$ if gene j is mutated in patient i , and $A_{ij} = 0$ otherwise.
- ▶ Goal: to identify a subset B of genes as **driver** gene.
- ▶ Vandin et al (2012, *Genome Res*) proposed two criteria:
 - ▶ Coverage: many patients with mutations in B , maximize $|\Gamma(B)|$ with $\Gamma(B) = \bigcup_{j \in B} \Gamma(j)$ and $\Gamma(j) = \{i : A_{ij} = 1\}$.
 - ▶ Exclusivity: mutations in B do not occur simultaneously on any patient, minimize $\omega(B) = \sum_{j \in B} |\Gamma(j)| - |\Gamma(B)|$.
- ▶ Overall, minimize

$$f(B) = \frac{\omega(B)}{n} - \frac{|\Gamma(B)|}{n} = \frac{1}{n} \sum_{j \in B} |\Gamma(j)| - \frac{2}{n} |\Gamma(B)|. \quad (1)$$

- ▶ Challenge: a combinatorial (i.e. NP-hard) problem!
not feasible to have an exact solution.
use approximate solutions, e.g. Monte Carlo methods...

Existing approaches

- ▶ Dendrix-MCMC: Vandin et al (2012, *Genome Res*);
- ▶ Multi-dendrix-MCMC: Leiserson et al (2013, *PLOS Comp Biol*);
- ▶ Binary linear programming (BLP) and genetic algorithm (GA): Zhao et al (2012, *Bioinformatics*);

New formulation

- ▶ Define: $B = B(\beta) = \{j \in V : |\beta_j| \neq 0\}$.
- ▶ Rewrite (1) as

$$f(B(\beta)) = \frac{1}{n} \sum_{j=1}^p I(|\beta_j| \neq 0) A_{\cdot j} - \frac{2}{n} \sum_{i=1}^n I\left(\sum_{j=1}^p A_{ij} I(|\beta_j| \neq 0) \neq 0\right).$$

- ▶ Challenge: discontinuous indicator function $I(\beta_j \neq 0)$.
- ▶ Our solution: use a TLP to approximate $I(\cdot)$:

$$\min(|\beta_j|/\tau_1, 1) \rightarrow I(|\beta_j| \neq 0) \quad \text{as } \tau_1 \rightarrow 0^+.$$

- ▶ Note: TLP as a non-convex penalty better than the popular Lasso in regularization, e.g. regression/classification, Gaussian graphical models (GGMs) (Shen et al 2012, *JASA*); fusion: network-based regression (Kim et al 2013, *Biometrics*; Zhu et al 2013, *JASA*), multiple GGMs (Gao et al 2016, *EJS*),

$$\tau_1 \min(|\beta_j|/\tau_1, 1) = \min(|\beta_j|, \tau_1).$$

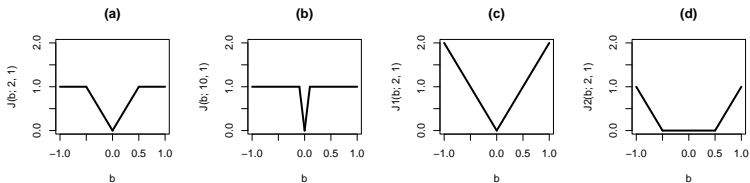


Figure: TLP $J_T(b; \tau)$ (a) with $\tau = 0.5$, or (b) with $\tau = 0.1$; J_T in (a) is decomposed into a difference of two convex functions J_1 in (c) and J_2 in (d).

Computation: DC algorithm

- ▶ New target:

$$S(\beta) = \frac{1}{n} \sum_{j=1}^p \min(\beta_j/\tau_1, 1) A_{\cdot j} - \frac{2}{n} \sum_{i=1}^n \min\left(\sum_{j=1}^p A_{ij} \beta_j/\tau_1, 1\right) + \lambda \sum_{j=1}^p \min(\beta_j/\tau_2, 1) + \frac{\alpha}{n} \sum_{j=1}^p \beta_j^2, \quad (3)$$

with respect to $\beta = (\beta_1, \dots, \beta_p)' \in [0, +\infty)^p$.

- ▶ Note: The TLP and the ridge penalty ensure sparse and proper solutions.
- ▶ DC decomposition of TLP:

$$\min\left(\frac{|z|}{\tau}, 1\right) = \frac{|z|}{\tau} - \max\left(\frac{|z|}{\tau} - 1, 0\right).$$

- ▶ At iteration m with the current estimate,

$$S^{(m)}(\beta) = \beta' \left(\text{diag}(\mathbf{A}.) I(\hat{\beta}^{(m-1)} \leq \tau_1) / n\tau_1 + \lambda I(\hat{\beta}^{(m-1)} \leq \tau_2) / \tau_2 - 2\mathbf{A}./n\tau_1 \right) + \frac{2}{n} \sum_{i=1}^n \max\left(\sum_{j=1}^p A_{ij} \beta_j / \tau_1 - 1, 0\right) + \frac{\alpha}{n} \beta' \beta, \quad (4)$$

- ▶ $S^{(m)}(\beta)$ is convex; can apply Matlab CVX or subgradient descent.
- ▶ Theorem: The above DC algorithm converges to a local minimizer monotonically in finite steps.
- ▶ Tuning parameter selection: $\alpha = 0.001$, $\tau_1 = 1$; others by CV.

Table: Applied to the mutation data of glioblastoma multiforme (TCGA 2008), the new method MCSS identified multiple sets of low-cost mutated genes, grouped in terms of associated pathways.

Pathway	Core mutations	\tilde{B}	$f(\tilde{B})$
p53 signalling	<i>CDKN2A, MDM2, MDM4, TP53</i>	<i>(CDKN2A, MDM2, MDM4, TP53)</i>	-55
		<i>(CDKN2A, DTX3, TP53)</i>	-57
		<i>(CDKN2A, TP53)</i>	-53
		<i>(CDKN2B, TP53)</i>	-53
RB signalling	<i>CDKN2A/B, CDK4, RB1</i>	<i>(CDKN2B, CYP27B1, RB1)</i>	-62
		<i>(CDKN2B, ERBB2, RB1, TSPAN31)</i>	-64
		<i>(CDKN2A, CYP27B1, RB1)</i>	-56
		<i>(CDKN2B, CYP27B1, NF1)</i>	-56
		<i>(CDKN2A, CYP27B1, NF1)</i>	-54
		<i>(CDKN2B, CYP27B1)</i>	-54
RAS signalling	<i>EGFR, NF1</i>	<i>(EGFR, KDR, NF1)</i>	-52
		<i>(MTAP, TP53, TSFM)</i>	-56
		<i>(CYP27B1, MTAP, PTEN)</i>	-55
		<i>(CDK4, MTAP, PTEN)</i>	-55
		<i>(EGFR, TP53)</i>	-52

Table: Results in Simulation I based on 100 simulation replications with $(p_1, p_2, p_3) = (0.95, 0.01, 0.05)$. The sample means (SD in parentheses) of correct (C) or incorrect (IC) numbers of non-zero estimates, average differences of the cost (ADC) between the true gene subset $B_0 = \{1, 2, 3, 4\}$ and the estimated subset \hat{B} , that is, $\frac{f(B_0) - f(\hat{B})}{n}$, and the running time (RT) (in minutes) of the algorithms.

n	p	Method	C	IC	ADC	\hat{c}_1 [c_1]	\hat{c}_2 [c_2]	RT
50	1000	MCSS	4 (0)	0 (0)	0 (0)	.95 [.95]	.01 [.00]	.22 (0)
		Dendrix	3.80 (.41)	.50 (.94)	-.02 (.04)	.94 [.95]	.01 [.00]	16.89 (0)
		Mdendrix	3.90 (.30)	.15 (.36)	-.01 (.03)	.95 [.95]	.01 [.00]	.81 (0)
		BLP	3.39 (.86)	3.82 (2.59)	.05 (.03)	.99 [.95]	.00 [.00]	.01 (0)
		GA	3.90 (.38)	2.13 (1.69)	.04 (.03)	.98 [.95]	.01 [.00]	2.97 (0)
50	10000	MCSS	4 (0)	0 (0)	0 (0)	.95 [.95]	.01 [.01]	1.67 (0)
		Dendrix	1.25 (1.02)	5.96 (3.62)	-.24 (.04)	.83 [.95]	.02 [.01]	67.06 (0)
		Mdendrix	1.45 (1.23)	5.25 (4.02)	-.24 (.03)	.83 [.95]	.03 [.00]	1.88 (0)
		BLP	3.42 (.91)	2.72 (2.06)	.05 (.03)	.99 [.95]	.01 [.00]	.27 (0)
		GA	3.93 (.25)	1.50 (.92)	.04 (.02)	.98 [.95]	.01 [.00]	284.92 (0)

Discussion

- ▶ A general method to approximate the indicator function; more applications (and modifications) may be worthwhile.
- ▶ Considered integrative analysis of mutation and gene expression data.
improved performance with GE data;
non-linear; binary linear programming (BLP) not applicable.
- ▶ TLP is a general non-convex penalty applied to many high-dimensional problems.
can be regarded as a refined version of Lasso; TLP=Lasso with larger τ ;
better empirical and theoretical performance;
computationally more demanding (with an extra tuning parameter τ).

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▶ **Thank you!**