

# Statistical Methods for Behavioral Economics Data in Tobacco Research

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# Outline

- 1 Background
- 2 Survival Analysis Methods Research
  - Recurrent event gap time methods
  - Intermittently observed time-dependent covariates
  - Composite endpoints with component-wise censoring
  - Graphical methods for survival endpoints in clinical trials
  - Collaborators
- 3 Statistical Methods for Behavioral Economics Data in Tobacco Research
  - Cigarette Purchase Task (CPT) Data
  - Left-censored mixed effects model
  - Two-part mixed effects model
  - Bayesian hierarchical model for meta-analysis
  - Collaborators and reference

## 1 Background

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# Background

## Education:

- 1997 **B.S.**, Geography, Peking University, China  
**Minor**, Computer Science, Peking University, China
- 2000 **M.S.**, Geomorphology, Peking University, China
- 2005 **Ph.D.**, Biostatistics, Johns Hopkins Bloomberg School of Public Health, Baltimore, Maryland

## Employment:

- 2006 – present Biostatistics Core, Masonic Cancer Center, U of M
- 2006 – 2013 **Assistant Professor**, non-tenure track  
Biostatistics, School of Public Health, U of M
- 2013 – 2020 **Associate Professor**, non-tenure track  
Biostatistics, School of Public Health, U of M
- 2020 – present **Professor**, non-tenure track  
Biostatistics, School of Public Health, U of M

# Method Research

- Recurrent event, survival data, and correlated data
- Design and analysis of clinical trials
- Statistical research motivated by collaboration experience. For example,
  - recurrent infections in BMT patients
  - correlated data in cigarette purchase task
  - AUC of biomarkers for smokers

PI'ed grants for method research:

- NIH/NIMH R03, Statistical methods for bivariate alternating recurrent event data, 2017 - 2019
- NIH/NCI R03, Statistical methods for analyzing data of recurrent infections after hematopoietic cell transplantation, 2014 - 2016
- Grant-In-Aid, Quantile regression models for recurrent gap time data, 2012 - 2014
- Minnesota Medical Foundation, Statistical methods for measuring relative reinforcing efficacy using drug purchase task instrument data, 2012 - 2013

# Collaborative Research

## ● Tobacco and chemoprevention

### Ongoing:

- NIH/NCI P01, Consortium on Methods Evaluating Tobacco (COMET): filter ventilation and product standards (Hatsukami/Shields, Core C: Luo), 2017-2023
- FDA/NIDA U54, Evaluating New Nicotine Standards for Cigarettes (Donny/Hatsukami, Core C: Koopmeiners), 2017-2023
- NIH/NCI R01, Lung cancer prevention and treatment by targeting ALDH1 and CD44 expressing putative lung cancer stem cells (PI: Kassie), 2019-2024
- NIH/NIDA R01, Impact of sugars on tobacco product toxicity and abuse liability (Stepanov/Hatsukami), 2020-2023
- NIH/NIEHS R01, Biomarker phenotypes of air pollution and cancer risk in India (Stepanov/Dikshit), 2021-2026
- NIH R01, Evaluating cigarette relighting behavior: Prevalence, correlates, toxicant exposure, and implications for cessation (Steinberg/Stepanov/Heckman), 2022-2026 (funded)

### Completed (selected):

- NIH/NHLBI R01, Enhancing smoking cessation in homeless populations (Okuyemi/Pratt), 2014 - 2019
- NIH/NCMHD P60, University of Minnesota Center for Health Disparities Research, Engagement and Training (CeHDRET) (Ahluwalia, Proj 1: Thomas), 2009 - 2016
- NIH/NHLBI R01, Enhancing Quit and Win Contests to improve cessation among college smokers (Thomas), 2009 - 2014

# Collaborative Research (Continued)

## ● BMT and Immunology

### Ongoing:

- NIH/NCI P01 NK cells, their receptors and unrelated donor transplant (Miller, Core B: Le), 2021 - 2026
- NIH/NCI R35, Viral priming and targeting NK cells against solid tumor malignancies (Miller), 2015-2022
- DOD, Driving natural killer cell immunotherapy in the castration resistant prostate cancer setting with novel tri-specific killer engager molecules (Felices), 2020-2023
- DOD, Leveraging low oxygen environments for improved natural killer cell immunotherapy (Kennedy), 2021-2023

### Pending:

- NIH/NCI P01, Off-the-shelf immune effector cells for hematological malignancies (Wagner, Core B: Le), 2022-2027

# Mentoring

- Clinical/translational researchers:
  - Dr. Antonella Borgatti, NIH K01, 2014-2019
  - Dr. Lucie Turcotte, CTSI Pre-K, 2014-2016
  - Dr. Christen Ebens, CTSI KL2, 2019-present
  - Serve as CTSI-Ed Annual Poster Session Judge: 2016, 2017, 2021
- PhD dissertations completed (or expected):
  - Chihyun Lee (2014)
  - Tianmeng Lyu (2018)
  - Anne Eaton (2020)
  - Sandra Castro-Pearson (2022 expected, advised jointly with Dr. Le)
- MS theses advised: 12 completed



# Service (MCC and University)

- MCC
  - Cancer Protocol Review Committee (CPRC): 2009-2016, 2018-present
  - Data Safety and Monitoring Committee (DSMC): 2016-2018
- CTSI/U of M
  - CTSI-Ed Program Annual Poster Session Judge: 2016, 2017, 2021
  - Undergraduate Research Symposium Judge: 2018
- School of Public Health and Biostatistics Division (current)
  - Faculty Consultative Committee (FCC), At-Large Representative, 2020-2023
  - Faculty Salary Equity Review Committee (SERC), Co-Chair, 2022-2024
  - Diversity, Climate, and Inclusion (DCI) Committee, Co-Chair, 2019-present

## Service (to the Profession)

- American Statistical Association (ASA) Twin Cities Chapter Representative to the Council of Chapters; 2018 - 2021.
- NIH Grant Reviews: Special Emphasis Panel, Tobacco Control Regulatory Research, Tobacco Regulatory Science A & Basic Science, Health Services Organization, Secondary Analyses of Existing Datasets of Tobacco Use and Health, Quality and Effectiveness; 2013-present.
- Journal Editorial Board: BMC Medicine (2015-2020), BMC Medical Research Methodology (2016-present)
- Journal Referee: 20+ journals
- DSMB for projects in other institutions: U of Iowa, U of Arizona

## 1 Background

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# Statistical Analysis Methods Research Overview

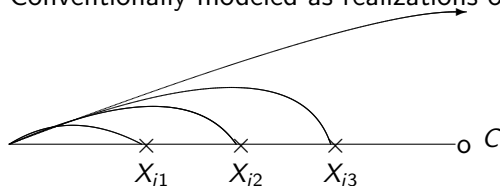
- Recurrent event gap time methods
  - Weighted risk-set & modified within-cluster resampling (i.e. “outputation”)
  - Induced smoothing (for recurrent gap time AFTM)
- Intermittently observed time-dependent covariates in recurrent events models
  - Kernel smoothing
- Composite endpoints with component-wise censoring
  - Kernel smoothing (on a marker process)
- Graphical methods for survival endpoints in clinical trials
  - An “ROC” approach

# Recurrent Event Gap Time Methods

Two time scales:

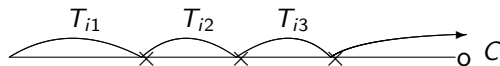
- Time since entering the study

Conventionally modeled as realizations of counting processes.



- Time since the last event

Focus on distribution of the gap time between consecutive events.



# Recurrent Event Gap Time Methods

## Tricky data structure

- Gap times are ordered and correlated
- Number of recurrent events is informative.
- Length bias
- Induced informative censoring
- Hence, clustered survival data methods (even with informative cluster size) cannot be directly applied to recurrent gap time data

# Recurrent Event Gap Time Methods - WCR

## Within-Cluster Resampling Method (WCR)

- In each resampling, one observation is randomly selected from each cluster to form a subsample of independent observations
- The resampling procedure is repeated a large number of times
- An estimator can be obtained through averaging over the estimates obtained from the resampled data.
- The variance of the WCR estimate can be estimated by the average of a consistent estimator of variance from each resampled dataset minus the empirical variance-covariance matrix of parameter estimates

Hoffman, Sen, and Weinberg (Biometrika, 2001), Follmann, Proschan, and Leifer (Biometrics, 2003), Williamson, Kim, Manatunga, and Addiss (SIM, 2008)

# Recurrent Event Gap Time Methods - WCR

A direct application of WCR method to recurrent gap time data?

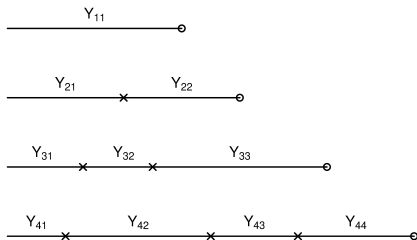
- Length bias
- Induced informative censoring

A **modified WCR method** for recurrent gap time data analysis under the conditions:

- Each individual recurrent event process is a renewal process
- The censoring time is independent of  $\{T_{i1}, T_{i2}, \dots\}$



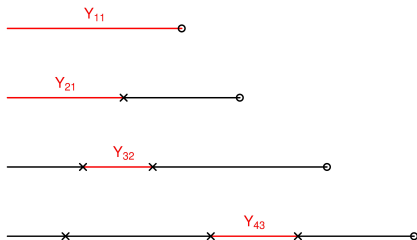
# Recurrent Event Gap Time Methods - WCR



$$m_i : \sum_{j=1}^{m_i-1} T_{ij} \leq C_i \text{ and } \sum_{j=1}^{m_i} T_{ij} > C_i$$

$$m_i^* = \max\{1, m_i - 1\}$$

# Recurrent Event Gap Time Methods - WCR



One gap time is randomly selected from each subject to form a subsample of independent observations

Then apply Kaplan-Meier estimator, Cox model, AFT model, etc, and repeat many times.

# Recurrent Event Gap Time Methods - WCR

- For each resampling:
  - $Y_{i1}$  is always chosen if subject  $i$  has only one censored gap time
  - Only uncensored gap times are selected if a subject has  $\geq 1$  events
  - The subsample consists of independent observations. Standard software can be readily applied.
- The WCR estimator can be obtained through averaging over the estimates obtained from the resampled data.

$$\hat{\beta}^{wcr} = B^{-1} \sum_{b=1}^B \hat{\beta}_b$$

$$\hat{\Sigma}^{wcr} = \frac{1}{B} \sum_{b=1}^B \hat{\Sigma}_b - \frac{1}{B-1} \sum_{b=1}^B (\hat{\beta}_b - \hat{\beta}^{wcr}) \otimes^2,$$

Hoffman, Sen and Weinberg (Bmka 2001)

# Recurrent Event Gap Time Methods - WRS

## The Weighted Risk Set (WRS) Method

The WCR method can be used to extend univariate survival analysis methods to analyze recurrent gap time data using the idea of [inverse weighting](#).

- Assign each observation in a risk set a weight that is proportional to the inverse of the "effective" cluster size,  $1/m_i^*$
- Weighted logrank test, AFTM, model checking, etc.

## Recurrent Event Gap Time Methods - Reference

**Luo X**, Huang C-Y. (2011). Analysis of recurrent gap time data using the weighted risk set method and the modified within-cluster resampling method. *Statistics in Medicine*, 30(4):301–311.

Huang C-Y, **Luo X**, Follmann D. (2011). A model checking method for the proportional hazards model with recurrent gap time data. *Biostatistics*, 12(3):535–547.

**Luo X**, Huang C-Y, Wang L. (2013). Quantile regression for recurrent gap time data. *Biometrics*, 69(2):375–385.

Lee CH, **Luo X\***, Huang C-Y, DeFor T, Brunstein CG, Weisdorf DJ. (2016). Nonparametric methods for analyzing recurrent gap time data with application to infections after hematopoietic cell transplant. *Biometrics*, 72(2):535–545.

Lee CH, Huang C-Y, Xu G, **Luo X**. (2018). Semiparametric regression model for alternating recurrent event data. *Statistics in Medicine*, 37(6):996–1008.

Lyu T, **Luo X\***, Xu G, Huang C-Y. (2018). Induced smoothing for rank-based regression with recurrent gap time data. *Statistics in Medicine*, 37(7):1086–1100.

Lee CH, Huang C-Y, DeFor T, Brunstein CG, Weisdorf DJ, **Luo X**. (2019). Semiparametric regression model for recurrent bacterial infections after hematopoietic stem cell transplantation. *Statistica Sinica*, 2019 July; 29(3):1489-1509.

# Intermittently Observed Time-Dependent Covariates

Now, focus on time-to-recurrent events data

Time-dependent covariates  $X(t)$ :

- Covariate values changing with time  $t$ .
- Usually measured at discrete time points or visits.
- Examples:
  - self-reported nicotine dependence of smokers (FTND) at monthly visits, related to the cessation outcome
  - bacterial infection status measured at regular visits, related to the risk of pharyngitis
  - Biomarkers, e.g., CD4 cell counts, related to the risk of opportunistic infections

# Intermittently Observed Time-Dependent Covariates

Notation:

- $N^*(t)$ : number of recurrent events occurring at or prior to time  $t$ .
- *Intensity function* of the recurrent event process  $N^*(t)$ :

$$\lambda(t|H(t)) = \lim_{\Delta \rightarrow 0^+} \frac{P(N^*(t + \Delta) - N^*(t) > 0 | H(t))}{\Delta} \quad (\text{Conditional})$$

where  $H(t)$  represents the event history at or prior to time  $t$ .

- *Rate function* of the recurrent event process  $N^*(t)$ :

$$\lambda(t) = \lim_{\Delta \rightarrow 0^+} \frac{P(N^*(t + \Delta) - N^*(t) > 0)}{\Delta} \quad (\text{Marginal})$$

# Intermittently Observed Time-Dependent Covariates

Proportional rates model (Lin et al., 2000):

$$\lambda\{t|\mathbf{Z}_i(t)\} = \lambda_0(t) \exp\{\boldsymbol{\beta}^\top \mathbf{Z}_i(t)\}.$$

Additive rates model (Schaubel et al., 2006):

$$\lambda\{t|\mathbf{Z}_i(t)\} = \lambda_0(t) + \boldsymbol{\beta}^\top \mathbf{Z}_i(t).$$

Additive-multiplicative rates model (Liu et al., 2010):

$$\lambda\{t|\mathbf{W}_i(t)\} = \boldsymbol{\gamma}^\top \mathbf{Z}_i(t) + \exp\{\boldsymbol{\beta}^\top \mathbf{X}_i(t)\} \lambda_0(t),$$

- They all allow time-dependent covariates theoretically.
- Their model estimation procedures require the time-dependent covariates to be continuously observed throughout the entire follow up period for each subject.



# Intermittently Observed Time-Dependent Covariates

Proportional rates model (Lin et al., 2000):

$$U(\beta) = n^{-1} \sum_{i=1}^n \int_0^{\tau} \left\{ \mathbf{Z}_i(t) - \frac{S^{(1)}(t, \beta)}{S^{(0)}(t, \beta)} \right\} dN_i(t),$$

where  $S^{(k)}(t) = n^{-1} \sum_{l=1}^n Y_l(t) \mathbf{Z}_l(t)^{\otimes k} \exp\{\beta^T \mathbf{Z}_l(t)\}$ ,  
 $N_i(t) = N_i^*(\min\{t, C_i\})$ ,  $Y_i(t) = I(C_i \geq t)$ .

Additive rates model (Schaubel et al., 2006):

$$U(\beta) = n^{-1} \sum_{i=1}^n \int_0^{\tau} \{\mathbf{Z}_i(t) - \bar{\mathbf{Z}}(t)\} dN_i(t) - \left[ n^{-1} \sum_{i=1}^n \int_0^{\tau} \{\mathbf{Z}_i(t) - \bar{\mathbf{Z}}(t)\}^{\otimes 2} Y_i(t) dt \right] \beta,$$

where  $\bar{\mathbf{Z}}(t) = \frac{n^{-1} \sum_{i=1}^n Y_i(t) \mathbf{Z}_i(t)}{n^{-1} \sum_{i=1}^n Y_i(t)}$ .

# Intermittently Observed Time-Dependent Covariates

Additive-multiplicative rates model (Liu et al., 2010):

$$U(\boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^n \int_0^{\tau} \{\mathbf{D}_i(\boldsymbol{\theta}, u) - \bar{\mathbf{D}}(\boldsymbol{\theta}, u)\} \{dN_i(u) - Y_i(u) \boldsymbol{\gamma}^T \mathbf{Z}_i(u) du\}.$$

where

$$\mathbf{D}_i(\boldsymbol{\theta}, t) = \begin{pmatrix} \mathbf{Z}_i(t) / \exp\{\boldsymbol{\beta}^T \mathbf{X}_i(t)\} \\ \mathbf{X}_i(t) \end{pmatrix},$$

$$\bar{\mathbf{D}}(\boldsymbol{\theta}, t) = \frac{\frac{1}{n} \sum_{i=1}^n Y_i(t) \mathbf{D}_i(\boldsymbol{\theta}, t) \exp\{\boldsymbol{\beta}^T \mathbf{X}_i(t)\}}{\frac{1}{n} \sum_{i=1}^n Y_i(t) \exp\{\boldsymbol{\beta}^T \mathbf{X}_i(t)\}}.$$

# Intermittently Observed Time-Dependent Covariates

Additive Rates Model with Intermittently Observed Time-Dependent Covariates (Lyu et al., 2021)

By solving  $U(\beta) = 0$ , the estimator has a **closed-form** expression:

$$\begin{aligned} \hat{\beta} &= \left[ n^{-1} \sum_{i=1}^n \int_0^{\tau} \{\mathbf{Z}_i(t) - \bar{\mathbf{Z}}(t)\}^{\otimes 2} Y_i(t) dt \right]^{-1} \left[ n^{-1} \sum_{i=1}^n \int_0^{\tau} \{\mathbf{Z}_i(t) - \bar{\mathbf{Z}}(t)\} dN_i(t) \right] \\ &= \left[ \int_0^{\tau} \left( n^{-1} \sum_{i=1}^n Y_i(t) \frac{S^{(2)}(t)}{S^{(0)}(t)} - n^{-1} \sum_{i=1}^n Y_i(t) \left\{ \frac{S^{(1)}(t)}{S^{(0)}(t)} \right\}^{\otimes 2} \right) dt \right]^{-1} \\ &\quad \left[ n^{-1} \sum_{i=1}^n \int_0^{\tau} \mathbf{Z}_i(t) dN_i(t) - \int_0^{\tau} \frac{S^{(1)}(t)}{S^{(0)}(t)} \left\{ n^{-1} \sum_{i=1}^n dN_i(t) \right\} \right]. \end{aligned}$$

Replace  $\frac{S^{(k)}(t)}{S^{(0)}(t)}$  with  $\frac{\widehat{S}_h^{(k)}(t)}{\widehat{S}_h^{(0)}(t)}$ :

$$\frac{\widehat{S}_h^{(k)}(t)}{\widehat{S}_h^{(0)}(t)} = \frac{n^{-1} \sum_{i=1}^n \int_0^{\tau} K_h(t-u) Y_i(u) \mathbf{Z}_i(u)^{\otimes k} dO_i(u)}{n^{-1} \sum_{i=1}^n \int_0^{\tau} K_h(t-u) Y_i(u) dO_i(u)}$$

## Kernel Smoothing Methods - Reference

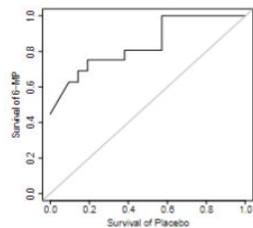
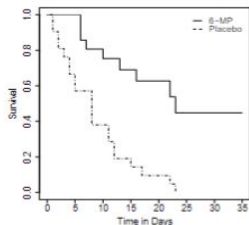
Lyu T, **Luo X**, Huang C-Y, Sun Y. (2021). Additive rates model for recurrent event data with intermittently observed time-dependent covariates. *Statistical Methods in Medical Research*, 2021 October; 30(10):2239–2255.

Lyu T, **Luo X**, Sun Y. (2021). Additive-multiplicative rates model for recurrent event data with intermittently observed time-dependent covariates. *Journal of Data Science*, 2021 Nov 4; 19(4), 615–633.

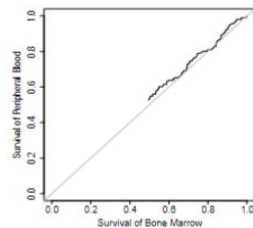
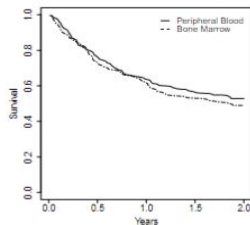
Eaton A, Sun Y, Neaton J, **Luo X**. (2022+). Nonparametric estimation in an illness-death model with component-wise censoring. *Biometrics*, 2021 Apr 29. doi: 10.1111/biom.13482. Online ahead of print.

# Graphical Methods for Survival Data in Clinical Trials

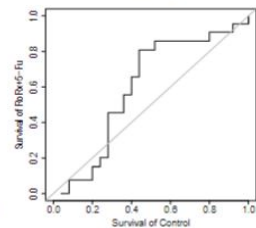
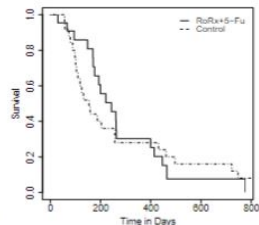
(a) 6-MP vs. Placebo Trial



(b) Peripheral Blood vs. Bone Marrow Trial



(c) Radiation and 5-Fluorouracil vs. Control Trial



# Graphical Methods for Survival Data - Reference

Castro-Pearson S, Le C, **Luo X**. (2022). Two-sample survival probability curves: A graphical approach for the analysis of time to event data in clinical trials. *Contemporary Clinical Trials*, 2022 Apr 1; 115:106707.

Castro-Pearson S, **Luo X**, Le C, et al. A graphical approach for the analysis of time to event data with competing risks in clinical trials. (Under preparation)

Castro-Pearson S, Le C, **Luo X**, et al. Monitoring randomization procedures in clinical trials. (Under preparation)

# Collaborators in Survival Analysis



## PhD Students



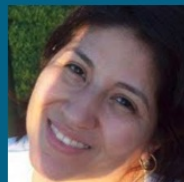
**Dr. Chi Hyun Lee**



**Dr. Tianmeng Lyu**



**Dr. Anne Eaton**



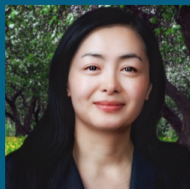
**Ms. Sandra Castro  
Pearson**

## Collaborators in Survival Analysis



**Dr. Chiung-Yu Huang**

UCSF



**Dr. Lan Wang**

University of Miami



**Dr. Gongjun Xu**

University of Michigan



**Dr. Yifei Sun**

Columbia University



**Dr. Chap T. Le**

University of Minnesota



# Statistical Methods for Behavioral Economics Data in Tobacco Research

## Overview

- Family Smoking Prevention and Tobacco Control Act (2009) and tobacco regulatory science
- Product and liability and relative reinforcing efficacy (RRE) and Cigarette Purchase Task (CPT) data
- Left-censored mixed effects model
- Two-part mixed effects model
- Bayesian hierarchical model for meta-analysis

# Family Smoking Prevention and Tobacco Control Act and Tobacco Regulatory Science

**Family Smoking Prevention and Tobacco Control Act** is a federal statute signed into law in June 22, 2009. The Act gives the **Food and Drug Administration (FDA)** the power to regulate the tobacco industry.

- Goal: discouraging minors and young adults from smoking
- New warnings and labels on packaging and advertisements
- Bans flavored cigarettes, places limits on the advertising of products to minors
- Requires tobacco companies to seek FDA approval for new products

# Family Smoking Prevention and Tobacco Control Act and Tobacco Regulatory Science

After the Family Smoking and Prevention and Tobacco Control Act was passed, tobacco research has growing even stronger.

- A new research territory is called [tobacco regulatory science](#)
- One of the fast growing research areas is studying the [Relative Reinforcing Efficacy \(RRE\)](#) (a central concept in psychopharmacology) of tobacco products
- The [drug purchase task](#) (e.g., cigarette purchase task/CPT) is a frequently used instrument for measuring RRE of a product.
- As part of those efforts, my colleagues and I have focused on the modelling and data analysis of CPT data.

# Statistical Methods for Behavioral Economics Data in Tobacco Research

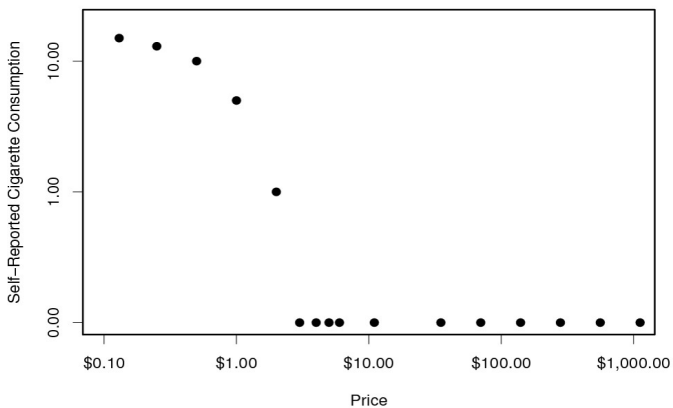


# Cigarette Purchase Task (CPT) Data

Imagine a TYPICAL DAY during which you smoke. The following questions ask how many cigarettes you would consume if they cost various amounts of money. Assume the following:

- Available cigarettes are your favorite brand
- You have the same income/savings that you have now
- You have **NO ACCESS** to any cigarettes or nicotine products other than those offered at these prices
- You would consume cigarettes that you request on that day (in other words, no stockpiling)

- |   |  |     |
|---|--|-----|
| 1. How many cigarettes would you smoke per day if they were each <u>free</u> ?    |  | [ ] |
| 2. How many cigarettes would you smoke per day if they were 1¢ each?              |  | [ ] |
| 3. How many cigarettes would you smoke per day if they were 5¢ each?              |  | [ ] |
| ...   |  |     |
| 18. How many cigarettes would you smoke per day if they were <u>\$560</u> each?   |  | [ ] |
| 19. How many cigarettes would you smoke per day if they were <u>\$1,120</u> each? |  | [ ] |



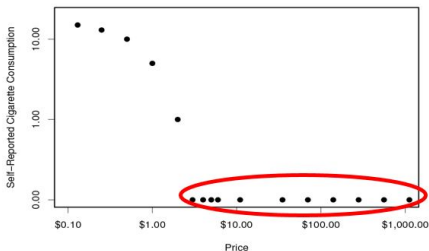
# Hursh and Silberberg's exponential demand curve

$$\log Q(p) = \log Q_0 + k(e^{-\alpha p} - 1),$$

- $p$ : price
- $Q$ : demand for a commodity (e.g. cigarettes)
- $Q_0$ : demand at price 0 (*intensity*)
- $k$ : range of log-consumption ( $= \log Q_0 - \log Q_\infty$ )
- *elasticity*( $p$ )  $= \partial \log Q(p) / \partial \log p$
- $Omax = \max\{Q(p) \cdot p\}$ : maximum expenditure
- $Pmax$ : price at which  $Omax$  is achieved
- **Breakpoint**: the first price for which the consumption is zero

# Hursh and Silberberg's exponential demand curve (continued)

- In application: for each subject's data  $(Q_{ij}, p_j), j = 1, 2, \dots$ , fit a nonlinear regression model to get one set of estimated parameters for each subject (overparameterized).
- How to deal with zero demand (when  $p \geq \text{breakpoint}$ )? The log of zero is negative infinity. Ignoring zero demand or arbitrarily imputing a small value will cause bias.





# Left-Censored Mixed Effects Model (Liao et al., 2013)

Goal: Fit the exponential demand curve to the data without having to delete or impute data

Assumption: the self-reported zero consumption at and beyond the breakpoint is a small nonzero consumption amount below a certain threshold ( $\omega$ ) that smokers do not bother to report - a limit of detection (LOD) or **left censoring** problem!

Model: A nonlinear mixed effects model

$$\log Q_{ij} = \mu_{ij} + \epsilon_{ij} = \log Q_{0i} + k(e^{-\alpha_i P_j} - 1) + \epsilon_{ij},$$

where

- $(\log Q_{0i}, \alpha_i)'$  is the vector of random effect (MVN with mean  $(\mu_0, \mu_\alpha)'$  and variances  $(\sigma_0^2, \sigma_\alpha^2)$  and covariance  $\rho\sigma_0\sigma_\alpha$ )
- $\epsilon_{ij}$  is the measurement error (MVN with mean 0 and variance  $\sigma_e^2$ ).

## Left-Censored Mixed Effects Model (continued)

The likelihood function for the left-censored mixed effects model has the same form as that for a correlated survival data with Type I left censoring and log-normal survival time (Klein & Moeschberger, 2003).

$$\prod_{i=1}^N \int \prod_{j=1}^{n_i} \left\{ \Phi \left( \frac{\log \omega - \mu_{ij}}{\sigma_e} \right) \right\}^{\delta_{ij}} \left\{ \sigma_e^{-1} \phi \left( \frac{\log Q_{ij} - \mu_{ij}}{\sigma_e} \right) \right\}^{1-\delta_{ij}} \phi(z_i) dz_i,$$

where

- $z_i = (\log Q_{0i} - \mu_0) / \sigma_0$
- $\delta_{ij}$  is the censoring indicator ( $=0$  if  $Q_{ij} \geq \omega$ ;  $=1$  if  $< \omega$ )
- $\Phi$  and  $\phi$  are cdf and pdf of standard normal, respectively

## Two-Part Mixed Effects Model (Zhao et al., 2016)

Motivation:  $\omega$  in the previous model is arbitrary;  $\delta_{ij}$  can be treated as a binary outcome, with a physical meaning (cessation), and can be modeled with a logistic regression

Part I model:

$$\text{logit}(\pi_{ij}) = \beta_0 + \beta_1 f(p_j) + a_i,$$

where  $\pi_{ij} = \Pr(\delta_{ij} = 1)$ , and  $f(\cdot)$  is a proper function of price.

Part II model for  $Q_{ij} > 0$  follows the same mixed effects model form as previously

$$\log Q_{ij} = \mu_{ij} + \epsilon_{ij} = (\log Q_0 + b_i) + k(e^{-(c+c_i)p_j} - 1) + \epsilon_{ij},$$

where  $\gamma_i = (a_i, b_i, c_i)'$  is the vector of random effects following a MVN with mean zero.

## Two-Part Mixed Effects Model (continued)

The likelihood function for two-part mixed effects model

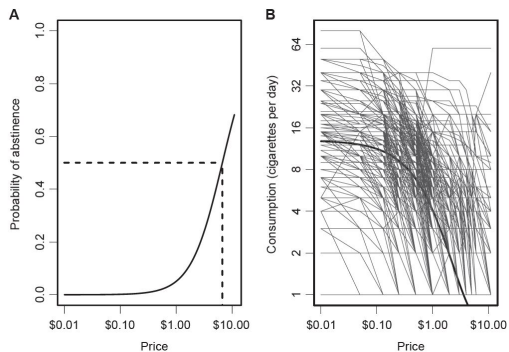
$$\prod_{i=1}^N \int \prod_{j=1}^{n_i} (\pi_{ij})^{\delta_{ij}} \left\{ (1 - \pi_{ij}) \sigma_e^{-1} \phi \left( \frac{\log Q_{ij} - \mu_{ij}}{\sigma_e} \right) \right\}^{1-\delta_{ij}} \phi(\gamma_i) d\gamma_i,$$

- Estimation can be carried out using the NLMIXED procedure in SAS.
- The estimated marginal parameters and the predicted random effects can be easily obtained to calculate each subject's intensity,  $O_{max}$ ,  $P_{max}$ , breakpoint,  $\alpha$  value, etc.

# Illustration with the Quit and Win Data (Thomas et al., 2016)

Quit and Win Study: 1,217 college smokers recruited from 19 two- or four-year colleges, studying two interventions (single vs. multiple contests and counseling vs. no counseling) on improving cessation

Data used: CPT data at baseline



## Illustration with the Quit and Win Data (continued)

Two demand parameters, intensity and breakpoint, significantly predicted reduction in cigarettes smoked per day (CPD) at 6 month among those who failed to quit.

*Percent reduction in cigarettes per day among smokers at 6 months*

Demand indices	$\beta$	95% CI	Type III test statistics	<i>p</i> -value
Intensity	-4.4%	(-5.4%, -3.3%)	71.40	<.0001
Omax	-0.2%	(-0.4%, 0.0%)	2.83	.09
Pmax	0.0%	(-3.7%, 3.9%)	0.00	.96
$\alpha$	-2.5%	(-34%, 29%)	0.02	.88
Breakpoint	-4.1%	(-5.6%, -2.6%)	29.17	<.0001

# Bayesian Hierarchical Model for Meta-Analysis (Zhang et al., in press)

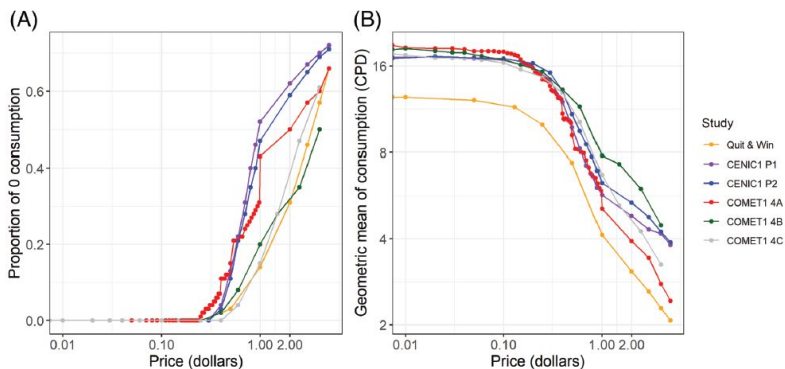
Six recent, multi-center tobacco studies coordinated by the University of Minnesota and designed and analyzed by the MCC Biostatistics Core.

Study	Full sample size	Analysis Sample size <sup>a,b</sup>	Age $\leq 25^c$
Quit & Win	1217	1214 (> 99%)	700 (58%)
CENIC1 P1	840	839 (> 99%)	137 (16%)
CENIC1 P2	1250	1227 (98%)	102 (8%)
COMET1 4A	224	179 (80%)	22 (12%)
COMET1 4B	211	181 (86%)	15 (8%)
COMET1 4C	295	223 (76%)	24 (11%)

Motivation: Extend the two-part mixed effects model by Zhao et al. by using a Bayesian hierarchical model to estimate the study-specific and population-averaged parameters simultaneously.

# Bayesian Hierarchical Model for Meta-Analysis (continued)

Summary of the observed data of each study



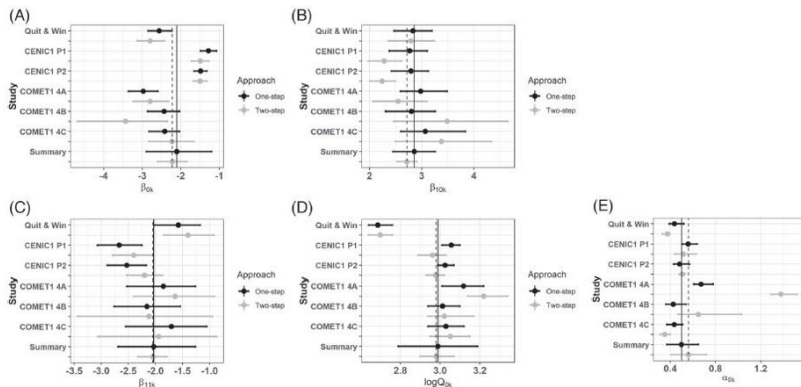


# Bayesian Hierarchical Model for Meta-Analysis (continued)

- Compared with the Zhao et al. model, the within-study model of the proposed Bayesian hierarchical model adopts a more flexible shape (splines) for the function of price,  $f(p)$  in the Part I model.
- Subject-level covariates are incorporated in the within-study model.
- Between-study model allows heterogeneity across studies, by assuming study-specific parameters to follow normal distributions.
- Noninformative priors were used.
- The likelihood and posterior distribution can be derived.
- MCMC was used for posterior computation, carried out using JAGS and the `rjags` package in R.

# Bayesian Hierarchical Model for Meta-Analysis (continued)

Compared with the two-step approach



- Shrinks toward the population-level parameters

# Bayesian Hierarchical Model for Meta-Analysis (continued)

- This work is meant to be a proof of concept for demonstrating the utility of the Bayesian framework for the analysis of complex data
- The proposed model for CPT data contains a substantial number of parameters, may require the study of its sensitivity to choices of priors
- The RRE indices which are not directly from the model parameters, e.g.  $O_{max}$  and  $P_{max}$ , can be easily obtained with posterior samples

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  - Dr. Warren K. Bickel (Virginia Tech)

  - Dr. Tracy T. Smith (University of South Carolina)

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